

Exercise 2.1-2.20 (Klein & Moeschberger, p.57-61)

2.1

X is from exponential distribution with $\lambda = 0.001$

$$(a) E(X) = \frac{1}{\lambda} = 1000$$

Mean of X is 1000.

$$(b) S(t_{0.5}) = \frac{1}{2}, \text{ where } S(\cdot) \text{ is survival function of } X.$$

$$\Rightarrow S(t_{0.5}) = \exp(-0.001 \times t_{0.5}) = \frac{1}{2}$$

$$\Rightarrow -0.001 \times t_{0.5} = \log\left(\frac{1}{2}\right)$$

$$\Rightarrow t_{0.5} = -1000 \log\left(\frac{1}{2}\right) \approx 301.03$$

\therefore median of X is 301.03

$$(c) \Pr(X > 2000) = S(2000) = \exp(-0.001 \times 2000) \approx 0.13534$$

2.2 $X \sim \text{Weibull} (\alpha = 2, \lambda = 0.001)$ the pdf is given by

$$f(x) = \alpha \lambda x^{\alpha-1} \exp(-\lambda x^\alpha), \quad x \geq 0$$

, and the survival function of X is

$$S(t) = \int_t^\infty \alpha \lambda x^{\alpha-1} \exp(-\lambda x^\alpha) dx = -\exp(-\lambda x^\alpha) \Big|_t^\infty = \exp(-\lambda t^\alpha)$$

$$(a) \Pr(X > 30) = \exp(-0.001 \times 30^2) \approx 0.40657$$

$$\Pr(X > 45) = \exp(-0.001 \times 45^2) \approx 0.13199$$

$$\Pr(X > 60) = \exp(-0.001 \times 60^2) \approx 0.02732$$

$$(b) E(X) = \int_0^\infty x \alpha \lambda x^{\alpha-1} \exp(-\lambda x^\alpha) dx$$

let $x^\alpha = u \quad du = \alpha x^{\alpha-1} dx$

$$= \int_0^\infty u^{\frac{1}{\alpha}} \lambda \exp(-\lambda u) du$$

$$= \lambda \frac{\Gamma\left(\frac{1}{\alpha} + 1\right)}{\lambda^{\frac{1}{\alpha}+1}} \int_0^\infty \frac{\lambda^{\frac{1}{\alpha}+1}}{\Gamma\left(\frac{1}{\alpha} + 1\right)} u^{\frac{1}{\alpha}+1-1} \exp(-\lambda u) du$$

$\therefore \frac{\lambda^{\frac{1}{\alpha}+1}}{\Gamma\left(\frac{1}{\alpha} + 1\right)} u^{\frac{1}{\alpha}+1-1} \exp(-\lambda u)$ is pdf of gamma distribution

$$\begin{aligned}
&= \lambda \frac{\Gamma\left(\frac{1}{\alpha} + 1\right)}{\lambda^{\frac{1}{\alpha}}} \times 1 = \frac{\Gamma\left(\frac{1}{\alpha} + 1\right)}{\lambda^{\frac{1}{\alpha}}} \\
\Rightarrow E(X) &= \frac{\Gamma\left(\frac{1}{2} + 1\right)}{0.001^{\frac{1}{2}}} = \frac{\frac{1}{2}\Gamma\left(\frac{1}{2}\right)}{0.001^{\frac{1}{2}}} = \frac{\frac{1}{2}\sqrt{\pi}}{0.001^{\frac{1}{2}}} \approx 28.02496
\end{aligned}$$

(c) The hazard function is that

$$\begin{aligned}
h(x) &= \frac{f(x)}{S(x)} = \frac{\alpha \lambda x^{\alpha-1} \exp(-\lambda x^\alpha)}{\exp(-\lambda x^\alpha)} = \alpha \lambda x^{\alpha-1} \\
\Rightarrow h(30) &= 2 \times 0.001 \times 30^{2-1} \approx 0.06 \\
h(45) &= 2 \times 0.001 \times 45^{2-1} \approx 0.09 \\
h(60) &= 2 \times 0.001 \times 60^{2-1} \approx 0.12
\end{aligned}$$

$$\begin{aligned}
(d) \quad S(t_{0.5}) &= \frac{1}{2} \\
\Rightarrow \exp(-0.001 \times t_{0.5}^2) &= \frac{1}{2} \\
\Rightarrow -0.001 \times t_{0.5}^2 &= \log(0.5) \\
\Rightarrow t_{0.5} &= \{-1000 \log(0.5)\}^{\frac{1}{2}} \approx 17.35022
\end{aligned}$$

\therefore median of X is 17.35022

2.3

$X \sim \text{log logistic}(\alpha = 1.5, \lambda = 0.01)$, the pdf is given by

$$\begin{aligned}
f(x) &= \frac{\alpha \lambda x^{\alpha-1}}{(1 + \lambda x^\alpha)^2}, x \geq 0 \\
(a) \quad S(t) &= \int_t^\infty \frac{\alpha \lambda x^{\alpha-1}}{(1 + \lambda x^\alpha)^2} dx = -\frac{1}{1 + \lambda x^\alpha} \Big|_t^\infty = \frac{1}{1 + \lambda t^\alpha} \\
\Rightarrow \Pr(X > 50) &= S(50) = \frac{1}{1 + 0.01 \times 50^{1.5}} \approx 0.22048 \\
\Pr(X > 100) &= S(100) = \frac{1}{1 + 0.01 \times 100^{1.5}} \approx 0.09091 \\
\Pr(X > 150) &= S(150) = \frac{1}{1 + 0.01 \times 150^{1.5}} \approx 0.05162
\end{aligned}$$

$$(b) \quad S(t_{0.5}) = \frac{1}{2}$$

$$\Rightarrow \frac{1}{1 + 0.01 \times t_{0.5}^{1.5}} = \frac{1}{2}$$

$$\Rightarrow 2 = 1 + 0.01 \times t_{0.5}^{1.5}$$

$$\Rightarrow t_{0.5} = 100^{\frac{1}{1.5}} = 21.54435$$

\therefore median of X is 21.54435

(c) The hazard function is that

$$h(x) = \frac{f(x)}{S(x)} = \frac{\alpha \lambda x^{\alpha-1}}{1 + \lambda x^\alpha}$$

$$h'(x) = \frac{\alpha(\alpha-1)\lambda x^{\alpha-2}(1 + \lambda x^\alpha) - \alpha \lambda x^{\alpha-1}(\alpha \lambda x^{\alpha-1})}{(1 + \lambda x^\alpha)^2}$$

$$= \frac{\alpha \lambda x^{\alpha-2} \{(\alpha-1)(1 + \lambda x^\alpha) - \alpha \lambda x^\alpha\}}{(1 + \lambda x^\alpha)^2}$$

$$= \frac{\alpha \lambda x^{\alpha-2} \{\alpha + \alpha \lambda x^\alpha - 1 - \lambda x^\alpha - \alpha \lambda x^\alpha\}}{(1 + \lambda x^\alpha)^2}$$

$$= \frac{\alpha \lambda x^{\alpha-2} \{\alpha - 1 - \lambda x^\alpha\}}{(1 + \lambda x^\alpha)^2}$$

let $h'(x) = 0$

$$\Rightarrow \frac{\alpha \lambda x^{\alpha-2} \{\alpha - 1 - \lambda x^\alpha\}}{(1 + \lambda x^\alpha)^2} = 0$$

$$\Rightarrow \{\alpha - 1 - \lambda x^\alpha\} = 0$$

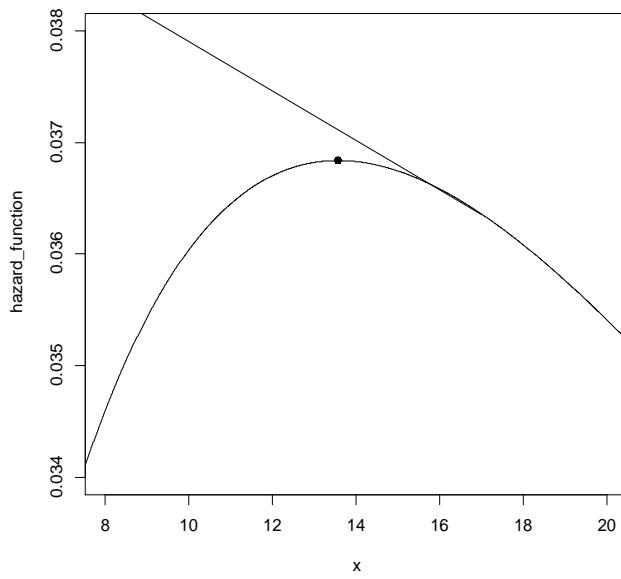
$$\Rightarrow x = \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}}$$

If $x > \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}}$ then $h'(x) < 0$

If $x < \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}}$ then $h'(x) > 0$

$h(x)$ is increasing when $0 \leq x < \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}} = \left(\frac{1.5 - 1}{0.01} \right)^{\frac{1}{1.5}} \approx 13.57209$

$h(x)$ is decreasing when $x > \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}} \approx 13.57209$



Black point is $x = \left(\frac{\alpha - 1}{\lambda} \right)^{\frac{1}{\alpha}} \approx 13.57209$, also the maximum of hazard function which

$$\lambda = 0.01, \alpha = 1.5.$$

(d)

$$E(X) = \int_0^\infty x \frac{\alpha \lambda x^{\alpha-1}}{(1 + \lambda x^\alpha)^2} dx$$

$$\text{Let } u = \frac{1}{1 + \lambda x^\alpha} \Leftrightarrow \left\{ \frac{1}{\lambda} (u^{-1} - 1) \right\}^\frac{1}{\alpha}$$

$$\Rightarrow du = -\frac{\alpha \lambda x^{\alpha-1}}{(1 + \lambda x^\alpha)^2} dx$$

$$\begin{aligned} &= \int_1^0 -\left\{ \frac{1}{\lambda} (u^{-1} - 1) \right\}^\frac{1}{\alpha} du = \lambda^{-\frac{1}{\alpha}} \int_0^1 \{u^{-1}(1-u)\}^\frac{1}{\alpha} du = \lambda^{-\frac{1}{\alpha}} \int_0^1 u^{-\frac{1}{\alpha}+1-1} (1-u)^{\frac{1}{\alpha}+1-1} du \\ &= \lambda^{-\frac{1}{\alpha}} B\left(-\frac{1}{\alpha} + 1, \frac{1}{\alpha} + 1\right), \text{ where } B(\cdot) \text{ is beta function} \end{aligned}$$

given that $\alpha = 1.5, \lambda = 0.01$

$$E(X) = 0.01^{-\frac{1}{1.5}} B\left(-\frac{1}{1.5} + 1, \frac{1}{1.5} + 1\right)$$

$$\approx 52.10283$$

2.4 X is a r.v ,with survival function that

$$S(x) = \exp[1 - \exp\{(\lambda x)^\alpha\}]$$

$$\begin{aligned} \Rightarrow f(x) &= -S'(x) = -\exp[1 - \exp\{(\lambda x)^\alpha\}][-\exp\{(\lambda x)^\alpha\}]\alpha(\lambda x)^{\alpha-1}\lambda \\ &= \exp\{(\lambda x)^\alpha\}\exp\{(\lambda x)^\alpha\}\alpha\lambda^\alpha x^{\alpha-1} \\ \Rightarrow h(x) &= \frac{f(x)}{S(x)} = \exp\{(\lambda x)^\alpha\}\alpha\lambda^\alpha x^{\alpha-1} \end{aligned}$$

(a) The hazard function is that

$$h(x) = \exp\{(\lambda x)^\alpha\}\alpha\lambda^\alpha x^{\alpha-1}$$

$$\begin{aligned} h'(x) &= \exp\{(\lambda x)^\alpha\}\alpha\lambda^\alpha x^{\alpha-1}\alpha\lambda^\alpha x^{\alpha-1} + \exp\{(\lambda x)^\alpha\}\alpha(\alpha-1)\lambda^\alpha x^{\alpha-2} \\ &= \exp\{(\lambda x)^\alpha\}\alpha(\alpha-1)\lambda^\alpha x^{\alpha-2}(\alpha\lambda^\alpha x^\alpha + \alpha - 1) \end{aligned}$$

Case I: given $\alpha = 0.5$

Let $h'(x) = 0$

$$\begin{aligned} \Rightarrow h'(x) &= \exp\{(\lambda x)^{0.5}\}0.5(0.5-1)\lambda^{0.5}x^{0.5-2}(0.5\lambda^{0.5}x^{0.5} + 0.5-1) = 0 \\ \Rightarrow (0.5\lambda^{0.5}x^{0.5} + 0.5-1) &= 0 \end{aligned}$$

$$\Rightarrow x = \frac{1}{\lambda}$$

If $x > \frac{1}{\lambda}$ then $h'(x) > 0$.

If $0 \leq x < \frac{1}{\lambda}$ then $h'(x) < 0$.

$h(x)$ is increasing when $x > \frac{1}{\lambda}$.

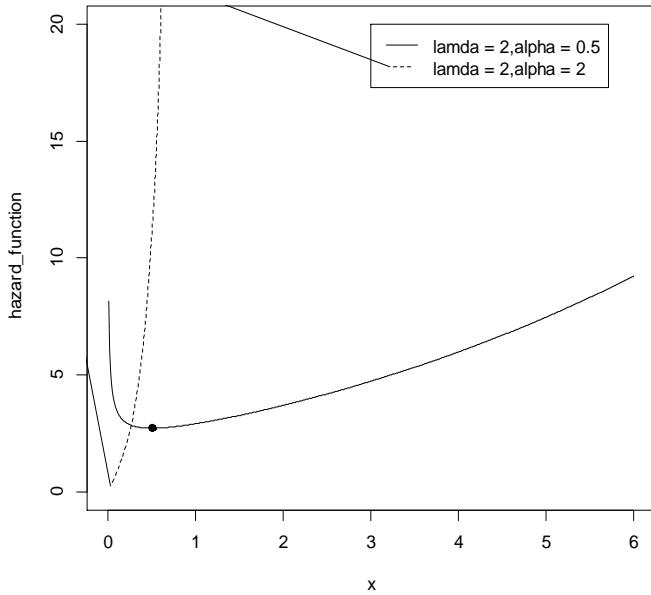
$h(x)$ is decreasing when $0 \leq x < \frac{1}{\lambda}$.

(b)

Case II: given $\alpha = 2$

$$h'(x) = \exp\{(\lambda x)^2\}2(2-1)\lambda^2x^{2-2}(2\lambda^2x^2 + 2-1) \geq 0, \forall x \geq 0$$

$h(x)$ is monotone increasing for all $x \geq 0$



Black point is $x = \frac{1}{\lambda} = \frac{1}{2}$, also the minimum of hazard function which

$\lambda = 2, \alpha = 0.5$. When $\lambda = 2, \alpha = 2$ of the hazard function is monotone increasing.

2.5

$X \sim \text{log normal}(\mu = 3.177, \sigma = 2.084)$

$\Rightarrow X = e^Y$, where $Y \sim N(\mu = 3.177, \sigma = 2.084)$

(a)

$E(X) = Ee^Y = M_Y(t)|_{t=1}$, with $M_Y(t)$ is mgf of normal

$$= \exp\left(\mu + \frac{1}{2}\sigma^2\right) \Big|_{\mu=3.177, \sigma=2.084} \approx 210.2985$$

$\Pr(X \leq t) = \Pr(\log X \leq \log t) = \Pr(Y \leq \log t) = \Phi\left(\frac{\log t - \mu}{\sigma}\right) \Big|_{\mu=3.177, \sigma=2.084}$, with $\Phi(\cdot)$ is cdf of $N(0,1)$.

$$\Rightarrow F(t_{0.5}) = \Phi\left(\frac{\log t_{0.5} - 3.177}{2.084}\right) = \frac{1}{2}$$

$$\Rightarrow \Phi^{-1}\left(\frac{1}{2}\right) = \frac{\log t_{0.5} - 3.177}{2.084}$$

$$\Rightarrow 3.177 = \log t_{0.5}$$

$$\Rightarrow t_{0.5} = \exp(3.177) \approx 23.97472$$

Mean of X is 210.2985 and median of X is 23.97472

(b)

$$S(100) = 1 - F(100) \approx 0.24658$$

$$S(200) = 1 - F(200) \approx 0.15436$$

$$S(300) = 1 - F(300) \approx 0.11267$$

(c)

$$h(x) = \frac{f(x)}{S(x)} = \frac{\frac{1}{2.084x} \phi\left(\frac{\log x - 3.177}{2.084}\right)}{1 - \Phi\left(\frac{\log x - 3.177}{2.084}\right)}, \text{ where } f(x) = F'(x) = \frac{1}{2.084x} \phi\left(\frac{\log x - 3.177}{2.084}\right),$$

with $\phi(\cdot)$ is pdf of $N(0,1)$

$$h'(x) = \frac{f'(x)S(x) - f(x)\{-f(x)\}}{\{S(x)\}^2} = \frac{f'(x)S(x) - f(x)\{-f(x)\}}{\{S(x)\}^2} = \frac{f'(x)}{S(x)} + \frac{\{f(x)\}^2}{\{S(x)\}^2},$$

$$\begin{aligned} \text{where } f'(x) &= \frac{-1}{2.084x^2} \phi\left(\frac{\log x - 3.177}{2.084}\right) + \frac{1}{2.084x} \phi\left(\frac{\log x - 3.177}{2.084}\right) \left(-\frac{\log x - 3.177}{2.084}\right) \frac{1}{2.084x} \\ &= \left(\frac{1}{2.084x}\right)^2 \phi\left(\frac{\log x - 3.177}{2.084}\right) \left\{ -2.084 - \left(\frac{\log x - 3.177}{2.084}\right) \right\} \\ &= \frac{1}{2.084x} \left\{ -2.084 - \left(\frac{\log x - 3.177}{2.084}\right) \right\} f(x) \\ h'(x) &= \frac{\frac{1}{2.084x} \left\{ -2.084 - \left(\frac{\log x - 3.177}{2.084}\right) \right\} f(x)}{S(x)} + \frac{\{f(x)\}^2}{\{S(x)\}^2} \end{aligned}$$

set $h'(x) = 0$

$$\begin{aligned} &\Rightarrow \frac{1}{2.084x} \left\{ -2.084 - \left(\frac{\log x - 3.177}{2.084}\right) \right\} + \frac{f(x)}{S(x)} = 0 \\ &\Rightarrow h(x) = \frac{1}{2.084x} \left\{ 2.084 + \left(\frac{\log x - 3.177}{2.084}\right) \right\} \\ &\Rightarrow x = h^{-1} \left[\frac{1}{2.084x} \left\{ 2.084 + \left(\frac{\log x - 3.177}{2.084}\right) \right\} \right] \end{aligned} \tag{1}$$

But x is hard to compute, I just set x_t is the solution of equation (1).

x_t is a critical point.

Set $h'(x) > 0$

$$\begin{aligned} &\Rightarrow \frac{1}{2.084x} \left\{ -2.084 - \left(\frac{\log x - 3.177}{2.084}\right) \right\} + \frac{f(x)}{S(x)} > 0 \\ &\Rightarrow h(x) < \frac{1}{2.084x} \left\{ 2.084 + \left(\frac{\log x - 3.177}{2.084}\right) \right\} \\ &\Rightarrow x < h^{-1} \left[\frac{1}{2.084x} \left\{ 2.084 + \left(\frac{\log x - 3.177}{2.084}\right) \right\} \right] \end{aligned}$$

$\Rightarrow h(x)$ is increasing when $0 < x < x_t$

Similarly $h'(x) < 0 \Rightarrow x > x_t$

By the previous results, we can know that $h(x)$ first increasing then decreasing change in the time x_t .

I use the bisection iteration to find the change point $x_t = 0.3470216$.

The following is the algorithm:

Step 1. Find two initial point x_L and x_R satisfy $h'(x_L)h'(x_R) < 0$

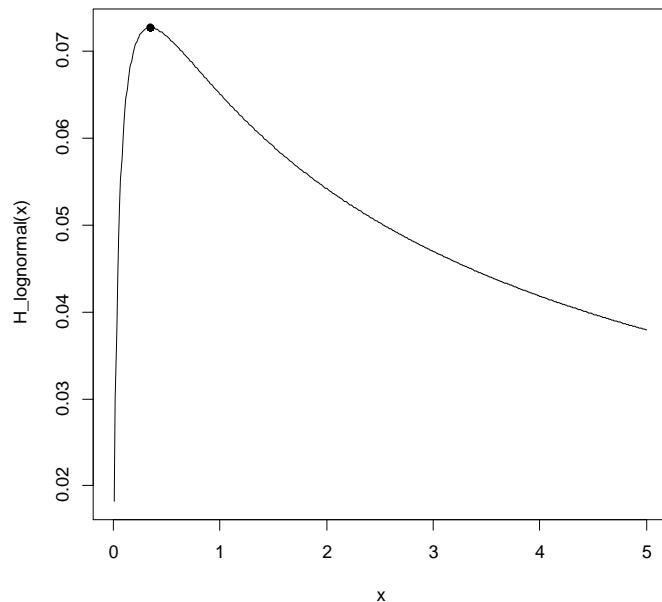
Step 2. Set $x_{New} = \frac{x_L + x_R}{2}$ if $h'(x_{New})h'(x_R) < 0$ then set $x_L = x_{New}$

otherwise set $x_R = x_{New}$

Step3.if $|x_L - x_R| < \varepsilon$, where ε is very small value then stop

otherwise goes to step1.

Graph1



Black point in the graph 1 is $x_t = 0.3470216$.

2.10

(a)

X is a r.v with hazard function

$$h(x) = \begin{cases} \theta_1 & , t_0 \leq x < t_1 \\ \theta_2 & , t_1 \leq x < t_2 \\ \theta_3 & , t_2 \leq x < t_3 \\ \vdots & \\ \theta_{k-1} & , t_{k-2} \leq x < t_{k-1} \\ \theta_k & , t_{k-1} \leq x \end{cases}$$

$$h(x) = \frac{f(x)}{S(x)} = \frac{d}{dx} \{-\log S(x)\}$$

for $t_{n-1} \leq x \leq t_n, 1 \leq n \leq k$

$$\begin{aligned} \int_{t_{n-1}}^{t_n} \frac{d}{dx} \{-\log S(x)\} dx &= \int_{t_{n-1}}^{t_n} \theta_n dx = \theta_n (t_n - t_{n-1}) \\ \Rightarrow -\log S(t_n) - \{-\log S(t_{n-1})\} &= \theta_n (t_n - t_{n-1}) \\ \Rightarrow -\log S(t_n) &= \theta_n (t_n - t_{n-1}) + \{-\log S(t_{n-1})\} \\ &= \theta_n (t_n - t_{n-1}) + \theta_{n-1} (t_{n-1} - t_{n-2}) + \{-\log S(t_{n-2})\} \\ &= \dots = \theta_n (t_n - t_{n-1}) + \theta_{n-1} (t_{n-1} - t_{n-2}) + \dots + \{-\log S(t_1)\} \\ &= \theta_n (t_n - t_{n-1}) + \theta_{n-1} (t_{n-1} - t_{n-2}) + \dots + \theta_1 (t_1 - t_0) + \{-\log S(t_0)\}, t_0 = 0 \\ &= \theta_n (t_n - t_{n-1}) + \theta_{n-1} (t_{n-1} - t_{n-2}) + \dots + \theta_1 t_1 \\ \Rightarrow S(t_n) &= \exp \left[-\{\theta_n (t_n - t_{n-1}) + \theta_{n-1} (t_{n-1} - t_{n-2}) + \dots + \theta_1 t_1\} \right] \\ \Rightarrow S(x) &= \begin{cases} \exp[-\{\theta_1 x\}], & t_0 \leq x < t_1 \\ \exp[-\{\theta_2 (x - t_1) + \theta_1 t_1\}], & t_1 \leq x < t_2 \\ \exp[-\{\theta_3 (x - t_2) + \theta_2 (t_2 - t_1) + \theta_1 t_1\}], & t_2 \leq x < t_3 \\ \vdots \\ \exp[-\{\theta_k (x - t_{k-1}) + \theta_{k-1} (t_{k-1} - t_{k-2}) + \dots + \theta_1 t_1\}], & t_{k-1} \leq x \end{cases} \end{aligned}$$

2.11

$$S(x) = \begin{cases} \exp \{-\lambda(x - \phi)^\alpha\}, & x \geq \phi \\ 1, & x < \phi \end{cases} \text{ is the survival function of } X.$$

(a)

$$f(x) = -S'(x) = \begin{cases} -\lambda \alpha (x - \phi)^{\alpha-1} \exp \{-\lambda(x - \phi)^\alpha\}, & x \geq \phi \\ 0, & x < \phi \end{cases}$$

$$h(x) = \frac{f(x)}{S(x)} = \begin{cases} -\lambda \alpha (x - \phi)^{\alpha-1}, & x \geq \phi \\ 0, & x < \phi \end{cases}$$

(b)

$$\Pr(X - \phi \geq t) = \Pr(X \geq t + \phi) = \exp(-\lambda t^\alpha)$$

$\therefore X - \phi \sim \text{Weibull}(\alpha, \lambda)$

$$E(X) = E(X - \phi + \phi) = E(X - \phi) + \phi$$

$$= \frac{\Gamma\left(\frac{1}{\alpha} + 1\right)}{\lambda^{\frac{1}{\alpha}}} + \phi \quad (\text{by 2.2(b)})$$

Given that $\alpha = 1, \lambda = 0.0075$ and $\phi = 100$

$$E(X) = \frac{\Gamma\left(\frac{1}{1} + 1\right)}{0.0075^{\frac{1}{1}}} + 100 \approx 233.333$$

by 2.2(d)

$Y \sim \text{Weibull}(\alpha, \lambda)$

$$\text{median of } Y \text{ is } \left(-\frac{1}{\lambda} \log \frac{1}{2}\right)^{\frac{1}{\alpha}}$$

$$\Rightarrow \text{median of } X \text{ is } \left(-\frac{1}{\lambda} \log \frac{1}{2}\right)^{\frac{1}{\alpha}} + \phi$$

given that $\alpha = 1, \lambda = 0.0075$ and $\phi = 100$

$$\text{median of } X \text{ is } \left(-\frac{1}{0.0075} \log \frac{1}{2}\right)^{\frac{1}{1}} + 100 \approx 192.41962$$

2.19

There is a bivariate survival function :

$$S(x, y) = (1-x)(1-y)(1+0.5xy), 0 < x < 1, 0 < y < 1$$

(a) The marginal survival function of X and Y are

$$S(x) = S(x, 0) = 1 - x$$

$$S(y) = S(0, y) = 1 - y$$

2.20

There is a bivariate survival function :

$$S(x, y) = \exp(-x - y - 0.5xy), 0 < x, 0 < y$$

(a) The marginal survival function of X and Y are

$$S(x) = S(x, 0) = \exp(-x)$$

$$S(y) = S(0, y) = \exp(-y)$$

- HW for Newton Raphson

$$l(\lambda) = \log L(\lambda) = \sum_{i=1}^n \log \{\exp(-\lambda l_i) - \exp(-\lambda r_i)\}$$

So the first order of loglikelihood is that

$$l'(\lambda) = \frac{\partial \log L(\lambda)}{\partial \lambda} = \sum_{i=1}^n \frac{-l_i \exp(-\lambda l_i) + r_i \exp(-\lambda r_i)}{\exp(-\lambda l_i) - \exp(-\lambda r_i)}$$

The scecond order of loglikelihood is that

$$\begin{aligned} l''(\lambda) &= \frac{\partial^2 \log L(\lambda)}{\partial \lambda^2} \\ &= \sum_{i=1}^n \frac{\{l_i^2 \exp(-\lambda l_i) - r_i^2 \exp(-\lambda r_i)\} \{\exp(-\lambda l_i) - \exp(-\lambda r_i)\} - \{-l_i \exp(-\lambda l_i) + r_i \exp(-\lambda r_i)\}^2}{\exp(-\lambda l_i) - \exp(-\lambda r_i)} \end{aligned}$$

Using Newton-Raphson technique to get MLE of λ

Newton-Raphson:

First choose a initial point λ_0 then updated λ by the following iteration

$$\lambda_{i+1} = \lambda_i - l'(\lambda_i) / l''(\lambda_i).$$

Stop when $|x_{i+1} - x_i| < \varepsilon$, where ε is very small value.

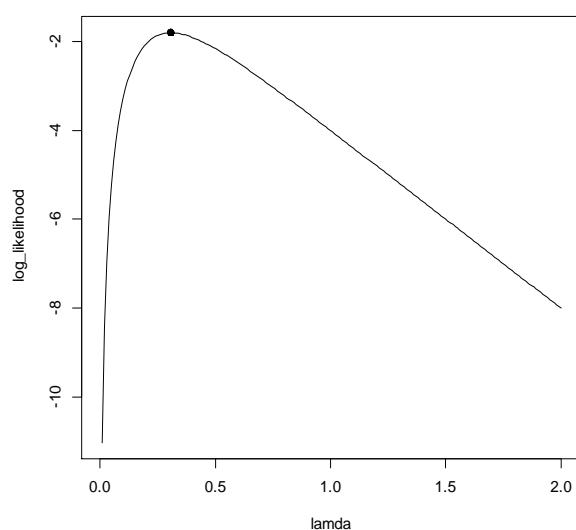
Given the data [0,7], [0,8], [0,5], [4,11]

$$\text{So that } l_1 = 0, l_2 = 0, l_3 = 0, l_4 = 4$$

$$r_1 = 7, r_2 = 8, r_3 = 5, r_4 = 11$$

Choose initial value $\lambda_0 = 0.2$ $\varepsilon = 10^{-5}$

$\hat{\lambda}_{MLE} = 0.3060561$ by Newton-Raphson technique.



Black point is $\hat{\lambda}_{MLE} = 0.3060561$.