

**Quiz 1, Statistical Inference I: Date 10/1 (2015 Fall)**

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1. Let  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\xi, \sigma^2)$  and  $X = (X_1, \dots, X_n)$ .

- 1) Show that the joint density of  $X = (X_1, \dots, X_n)$  forms an exponential family by specifying **all necessary components** (e.g., dimension  $s$ , natural parameters, parameter space, T functions, support, etc.).
- 2) Prove that the family is full rank.
- 3) Prove that the family is not full rank if  $\xi = \sigma$  is assumed.

Ans:

1)

$X_1, \dots, X_n \stackrel{iid}{\sim} N(\xi, \sigma^2)$ ,  $X = (X_1, \dots, X_n)$  let  $\theta = (\xi, \sigma^2)$ .

$$\begin{aligned} P_\theta(x) &= \prod_{i=1}^n f_i(x_i) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{\frac{-1}{2\sigma^2}(x_i - \xi)^2\right\} \\ &= \left(\frac{1}{2\pi\sigma^2}\right)^{\frac{n}{2}} \exp\left\{\frac{-1}{2\sigma^2} \sum_{i=1}^n (x_i - \xi)^2\right\} \\ &= \left(\frac{1}{2\pi}\right)^{\frac{n}{2}} \exp\left\{\frac{\xi}{\sigma^2} \sum_{i=1}^n x_i - \frac{1}{2\sigma^2} \sum_{i=1}^n x_i^2 - \frac{n}{2} \left(\frac{\xi^2}{\sigma^2} + \log \sigma^2\right)\right\}. \end{aligned}$$

Let

$$\eta_1(\theta) = \frac{\xi}{\sigma^2}, \quad \eta_2(\theta) = \frac{-1}{2\sigma^2}, \quad T_1(x) = \sum_{i=1}^n x_i, \quad T_2(x) = \sum_{i=1}^n x_i^2,$$

$$B(\theta) = \frac{n}{2} \left( \frac{\xi^2}{\sigma^2} + \log \sigma^2 \right) \text{ and } h(x) = \left( \frac{1}{2\pi} \right)^{\frac{n}{2}}.$$

Therefore,

$$P_\theta(x) = \exp \left\{ \sum_{i=1}^n \eta_i(\theta) T_i(x_i) - B(\theta) \right\} h(x)$$

is a two-dimensional exponential family with  $\chi = (-\infty, \infty)^n$  and  $\Theta = (-\infty, \infty) \times (-\infty, 0)$ .

2)

$$P_\theta(x) = \left(\frac{1}{2\pi}\right)^{\frac{n}{2}} \exp\left\{-\frac{\xi}{\sigma^2} \sum_{i=1}^n x_i - \frac{1}{2\sigma^2} \sum_{i=1}^n x_i^2 - \frac{n}{2} \left(\frac{\xi^2}{\sigma^2} + \log \sigma^2\right)\right\}.$$

$$\Theta = \left\{ \left( \frac{\xi}{2\sigma^2}, \frac{-1}{2\sigma^2} \right); \quad \xi \in \mathbf{R}, \sigma^2 > 0 \right\} = \{ (\eta_1, \eta_2); \quad \eta_1 \in \mathbf{R}, \eta_2 < 0 \}.$$

$\Theta$  contains a two-dimensional open rectangle hence (e.g., A(1, -1), B(1, -2), C(2, -1), D(2, -2) then ABCD is an two-dimensional open rectangle contained by  $\Theta$ ) hence it is full rank.

3)

Let  $\xi = \sigma^2$ .

$$P_\theta(x) = \left(\frac{1}{2\pi}\right)^{\frac{n}{2}} \exp\left\{-\frac{1}{\xi} \sum_{i=1}^n x_i - \frac{1}{2\xi} \sum_{i=1}^n x_i^2 - \frac{n}{2} (1 + \log \xi^2)\right\}.$$

$$\text{Since } \Theta = \left\{ \left( \frac{1}{\xi}, \frac{-1}{2\xi} \right); \quad \xi > 0 \right\} = \{ (\eta_1, \eta_2); \quad \eta_1 > 0, \eta_2 < 0 \}.$$

Therefore,  $\Theta$  does not contain a two-dimensional open rectangle hence it is not full rank.

**2.** Let  $X_1, X_2 \sim Poisson(\lambda)$  and  $X = (X_1, X_2)$ .

- 1) Derive the canonical form of the joint density of  $X = (X_1, X_2)$  by specifying all necessary components.

**Ans:**

$$\begin{aligned} P_\lambda(x) &= \frac{1}{x_1!} \lambda^{x_1} e^{-\lambda} \cdot \frac{1}{x_2!} \lambda^{x_2} e^{-\lambda} \\ &= \frac{1}{x_1! x_2!} \lambda^{x_1+x_2} e^{-2\lambda} \\ &= \frac{1}{x_1! x_2!} \exp \{ \log \lambda (x_1 + x_2) - 2\lambda \}. \end{aligned}$$

Let  $\eta(\lambda) = \log \lambda$ , we have

$$\begin{aligned} p(x|\eta) &= \frac{1}{x_1! x_2!} \exp \{ \eta(\lambda)(x_1 + x_2) - 2e^\eta \} \\ &= \exp \{ \eta(\lambda)T(x) - A(\eta) \} h(x) \end{aligned}$$

is the canonical form of a one-dimensional exponential family, where

$$\eta(\lambda) = \log \lambda, \quad T(x) = x_1 + x_2, \quad A(\eta) = 2e^\eta, \quad h(x) = \frac{1}{x_1! x_2!},$$

$$\chi = \{0, 1, \dots\} \times \{0, 1, \dots\} \text{ and } \Theta = (0, \infty).$$

2) Prove that the family is full rank.

**Ans:**

By 1)

$$p(x|\eta) = \exp \{ \eta(\lambda)T(x) - A(\eta) \} h(x).$$

Since  $\Theta = \{\eta; \eta > 0\}$  contains an one-dimensional open rectangle (e.g.,  $(1, 2)$ ) is an one-dimensional open rectangle contained by  $\Theta$ ) hence it is full rank.

3. Let  $p(x|\eta) = \exp\{\eta x - A(\eta)\}$ ,  $x \in X$ , be a density w.r.t. the Lebesgue measure  $\mu(x) = m(x)$ . Derive the form of  $A(\eta)$  and the natural parameter space  $\Theta$  when  $X = [0, \infty)$ .

**Ans:**

Since

$$p(x|\eta) = \exp\{\eta x - A(\eta)\}.$$

Hence

$$\begin{aligned} \int_X p(x|\eta) dx &= 1 \\ \Rightarrow \int_X \exp\{\eta x - A(\eta)\} dx &= 1 \\ \Rightarrow \int_X \exp\{\eta x\} dx &= e^{A(\eta)} < \infty. \end{aligned}$$

When  $X = [0, \infty)$ , consider two different cases of  $\eta > 0$  and  $\eta < 0$ .

$$\text{If } \eta > 0 \Rightarrow \int_0^\infty \exp\{\eta x\} dx = \infty.$$

$$\text{If } \eta < 0 \Rightarrow \int_0^\infty \exp\{\eta x\} dx < \infty$$

Therefore, the parameter space  $\Theta = (-\infty, 0)$ .

Then

$$\begin{aligned} \int_0^\infty \exp\{\eta x\} dx &= e^{A(\eta)} \\ \Rightarrow \frac{1}{\eta} e^{\eta x} \Big|_0^\infty &= e^{A(\eta)} \\ \Rightarrow 0 - \frac{1}{\eta} &= e^{A(\eta)} \\ \Rightarrow -\frac{1}{\eta} &= e^{A(\eta)} \\ \Rightarrow A(\eta) &= \log\left(-\frac{1}{\eta}\right). \end{aligned}$$

4. Let  $X = (X_1, \dots, X_n)$ , where  $X_j \stackrel{iid}{\sim} f_\theta(x_j) = \exp \left\{ \sum_{i=1}^s \eta'_i(\theta) T_i'(x_j) - B'(\theta) \right\} h'(x_j)$ ,

$j = 1, \dots, n$ . Show that the joint density of  $X = (X_1, \dots, X_n)$  forms an exponential family by specifying **all necessary components** (e.g., dimension, natural parameters, parameter space, T functions, support, etc.).

**Ans:**

$$X_j \stackrel{iid}{\sim} f_\theta(x_j) = \exp \left\{ \sum_{i=1}^s \eta'_i(\theta) T_i'(x_j) - B'(\theta) \right\} h'(x_j), \quad j = 1, \dots, n \quad \text{and} \quad X = (X_1, \dots, X_n).$$

$$\begin{aligned} P_\theta(x) &= \prod_{j=1}^n f_\theta(x_j) \\ &= \prod_{j=1}^n \exp \left\{ \sum_{i=1}^s \eta'_i(\theta) T_i'(x_j) - B'(\theta) \right\} h'(x_j) \\ &= \exp \{ \eta'_1(\theta) T_1'(x_1) + \eta'_2(\theta) T_2'(x_1) + \dots + \eta'_s(\theta) T_s'(x_1) \\ &\quad \eta'_1(\theta) T_1'(x_2) + \eta'_2(\theta) T_2'(x_2) + \dots + \eta'_s(\theta) T_s'(x_2) \\ &\quad \vdots \qquad \qquad \qquad \vdots \\ &\quad \eta'_1(\theta) T_1'(x_n) + \eta'_2(\theta) T_2'(x_n) + \dots + \eta'_s(\theta) T_s'(x_n) - nB'(\theta) \} \prod_{j=1}^n h'(x_j) \\ &= \exp \left\{ \eta'_1(\theta) \sum_{j=1}^n T_1'(x_j) + \eta'_2(\theta) \sum_{j=1}^n T_2'(x_j) + \dots + \eta'_s(\theta) \sum_{j=1}^n T_s'(x_j) - nB'(\theta) \right\} \prod_{j=1}^n h'(x_j) \\ &= \exp \left\{ \sum_{i=1}^s \left( \eta'_i(\theta) \sum_{j=1}^n T_i'(x_j) \right) - nB'(\theta) \right\} \prod_{j=1}^n h'(x_j) \\ &= \exp \left\{ \sum_{i=1}^s \eta_i(\theta) T_i(x) - B(\theta) \right\} h(x) \end{aligned}$$

is a s-dimensional exponential family with

$$\eta_i(\theta) = \eta'_i(\theta), \quad T_i(\theta) = \sum_{j=1}^s T_i'(x_j), \quad B(\theta) = nB'(\theta), \quad h(x) = \prod_{j=1}^n h'(x_j),$$

$$i = 1, 2, \dots, s \quad \Theta = \left\{ (\eta_1, \dots, \eta_s); \quad \int_{\chi} \exp \left\{ \sum_{i=1}^s \eta_i(\theta) T_i(x) \right\} h(x) < \infty \right\} \quad \text{with}$$

$$\text{support } \chi = \mathbb{R}^n.$$