

Quiz#1 High-dimensional Data Analysis, 2018 Spring

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Q [+6] Consider a simple linear model  $Y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$ ,  $\varepsilon_i \sim N(0, \sigma^2)$ .

+2 (1) [+2] Derive  $\hat{\beta}_0$  and  $\hat{\beta}_1$ . Assume the sample size is  $n$ , and  $n \geq 2$ .

residual sum of square  $RSS = \sum_{i=1}^n (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$

$\hat{\beta}_0, \hat{\beta}_1$  is the estimator that minimize the RSS, thus that satisfy the following equation:

$$\begin{cases} \frac{\partial}{\partial \hat{\beta}_0} RSS = \frac{\partial}{\partial \hat{\beta}_0} \sum_{i=1}^n (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2 = -2 \sum_{i=1}^n (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) \stackrel{\text{set}}{=} 0 \\ \frac{\partial}{\partial \hat{\beta}_1} RSS = \frac{\partial}{\partial \hat{\beta}_1} \sum_{i=1}^n (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2 = -2 \sum_{i=1}^n x_i (Y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) \stackrel{\text{set}}{=} 0 \end{cases} \Rightarrow \begin{cases} \bar{Y} - \hat{\beta}_0 - \hat{\beta}_1 \bar{x} = 0 \\ \sum_{i=1}^n x_i Y_i - \hat{\beta}_0 \sum_{i=1}^n x_i - \hat{\beta}_1 \sum_{i=1}^n x_i^2 = 0 \end{cases}$$

+1 (2) [+1] Derive  $SE(\hat{\beta}_0)$  (We should check the determinant of Hessian matrix to ensure that reach the global minimum, but here is ignored.)  $\Rightarrow \begin{cases} \hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{x} \\ \hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(Y_i - \bar{Y})}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{cases}$ , where  $\begin{cases} \bar{x} = \frac{\sum_{i=1}^n x_i}{n} \\ \bar{Y} = \frac{\sum_{i=1}^n Y_i}{n} \end{cases}$

Here use the result of part (3), first we show that

$$\begin{aligned} \text{Cov}(\bar{Y}, \hat{\beta}_1) &= \text{Cov}\left(\frac{1}{n} \sum_{i=1}^n Y_i, \frac{\sum_{j=1}^n (x_j - \bar{x}) Y_j}{\sum_{i=1}^n (x_i - \bar{x})^2}\right) = \frac{1}{n \sum_{i=1}^n (x_i - \bar{x})^2} \text{Cov}\left(\sum_{i=1}^n Y_i, \sum_{j=1}^n (x_j - \bar{x}) Y_j\right) \\ &= \frac{1}{n \sum_{i=1}^n (x_i - \bar{x})^2} \sum_{j=1}^n (x_j - \bar{x}) \sum_{i=1}^n \text{Cov}(Y_i, Y_j) \underbrace{\text{Cov}(Y_i, Y_j) = 0 \text{ if } i \neq j}_{=} = \frac{1}{n \sum_{i=1}^n (x_i - \bar{x})^2} \sum_{j=1}^n (x_j - \bar{x}) \cdot \text{Var}(Y_j) \\ &= \frac{\sigma^2}{n \sum_{i=1}^n (x_i - \bar{x})^2} \sum_{j=1}^n (x_j - \bar{x}) = 0 \end{aligned}$$

Thus,  $\text{Var}(\hat{\beta}_0) = \text{Var}(\bar{Y} - \hat{\beta}_1 \bar{x}) = \text{Var}(\bar{Y}) - 2\bar{x} \text{Cov}(\bar{Y}, \hat{\beta}_1) + \bar{x}^2 \text{Var}(\hat{\beta}_1)$

$$= \frac{\sigma^2}{n} - 0 + \bar{x}^2 \cdot \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \Rightarrow SE(\hat{\beta}_0) = \sqrt{\sigma^2 \left( \frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right)}$$

+1 (3) [+1] Derive  $SE(\hat{\beta}_1)$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(Y_i - \bar{Y})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n (x_i - \bar{x}) Y_i - \bar{Y} \sum_{i=1}^n (x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n (x_i - \bar{x}) Y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\begin{aligned} \Rightarrow \text{Var}(\hat{\beta}_1) &= \text{Var}\left(\frac{\sum_{i=1}^n (x_i - \bar{x}) Y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}\right) = \left(\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)^2 \text{Var}\left(\sum_{i=1}^n (x_i - \bar{x}) Y_i\right) \\ &= \left(\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)^2 \cdot \sum_{i=1}^n \text{Var}((x_i - \bar{x}) Y_i) = \left(\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)^2 \cdot \sum_{i=1}^n (x_i - \bar{x})^2 \text{Var}(Y_i) \\ &= \left(\frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)^2 \cdot \sum_{i=1}^n (x_i - \bar{x})^2 \cdot \text{Var}(\beta_0 + \beta_1 x_i + \varepsilon) = \frac{1}{\sum_{i=1}^n (x_i - \bar{x})^2} \cdot \text{Var}(\varepsilon_i) = \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{aligned}$$

+2 (4) [+2] Let  $Y = f(x) + \varepsilon$ ,  $\varepsilon \sim N(0, \sigma^2)$ , be a future observation. Write the formula of  $SE(\hat{\beta}_1) = \sqrt{\frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$

(prediction error) = (reducible error) + (irreducible error)

$$\begin{aligned} E[(Y - \hat{f}(x))^2] &= E[(f(x) + \varepsilon - \hat{f}(x))^2] \\ &= E[(f(x) - \hat{f}(x) + \varepsilon)^2] \quad (\text{where the notation } \hat{f}(x) \text{ is our prediction given the future observation } x) \\ &= E[(f(x) - \hat{f}(x))^2] + 2E[(f(x) - \hat{f}(x))\varepsilon] + E(\varepsilon^2) \\ &\stackrel{\text{indep}}{=} E[(f(x) - \hat{f}(x))^2] + 2E(f(x) - \hat{f}(x)) \cdot E(\varepsilon) + (E(\varepsilon^2) - [E(\varepsilon)]^2)) \\ &= E[(f(x) - \hat{f}(x))^2] + \text{Var}(\varepsilon) \end{aligned}$$

our simple linear regression  $E[(f(x) - \hat{\beta}_0 - \hat{\beta}_1 x)^2] + \sigma^2$

(prediction error) = (reducible error) + (irreducible error)