

HW #3 Quality Control, Spring 2013, Due 6/7 (Fri),

1. [+5] Perform profile monitoring for the data in Table 10.3. The value of α is chosen so that the in-control ARL is 200 [p.479 of the textbook]. Then, compare your chart with Fig. 10.28. [provide all necessary information, such as UCL/LCL, center, calculation of α , regression estimates, etc.; missing information leads to the reduction in your score]

2. [+2] I told you in class

3. [+6] Let $\mathbf{X}_{1k}, \dots, \mathbf{X}_{nk} \stackrel{iid}{\sim} N_3(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, $k = 1, \dots, m$, where $\boldsymbol{\mu} = \begin{pmatrix} -1 \\ -1 \\ -1 \end{pmatrix}$ and $\boldsymbol{\Sigma} = \begin{pmatrix} 1 & 0.5 & 0.5 \\ 0.5 & 1 & 0.5 \\ 0.5 & 0.5 & 1 \end{pmatrix}$.

Let $n = 25$, $m = 1000$, and $\alpha = 0.05$.

(1) [+1] Calculate $\boldsymbol{\Sigma}^{-1}$ (with derivation).

(2) [+2] Generate the data by R `mvrnorm` and draw a chi-square chart. Compare the observed ratio of out-of-control points with the nominal $\alpha = 0.05$.

(3) [+1] If $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ are unknown, we estimate them by $\bar{\bar{\mathbf{X}}} = \frac{1}{m} \sum_{k=1}^m \bar{\mathbf{X}}_k$ and $S = \frac{1}{m} \sum_{k=1}^m S_k$,

where $\bar{\mathbf{X}}_k = \frac{1}{n} \sum_{i=1}^n \mathbf{X}_{ik}$ and $S_k = \frac{1}{n-1} \sum_{i=1}^n (\mathbf{X}_{ik} - \bar{\mathbf{X}}_k)(\mathbf{X}_{ik} - \bar{\mathbf{X}}_k)'$. Compare $\bar{\bar{\mathbf{X}}}$ and S with $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$, respectively.

(4) [+2] Draw a chi-square chart by replacing $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ by $\bar{\bar{\mathbf{X}}}$ and S , respectively, and by using the same UCL and LCL as (2). Compare the observed ratio of out-of-control points with the nominal $\alpha = 0.05$. (This is called “the Hotelling T^2 chart”)

HW #3 Quality Control

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1.

We have $ARL = 200$, $\sigma = 0.06826$, $\bar{x} = \frac{0.76+3.29+8.89}{3} = 4.3133$ and

model $y_{ij} = 0.2817 + 0.9767x_j + \varepsilon_{ij}$.

```
> # data
> x = c(0.76,3.29,8.89)
> y = c(1.12,3.49,9.11,0.99,3.53,8.89,1.05,
+       3.46,9.02,0.76,3.75,9.3,0.96,3.53,
+       9.05,1.03,3.52,9.02)
>
> # center data
> x_bar = mean(x)
> centered_x = x - x_bar
>
> # model: yij = 0.2817 + 0.9767xj + error
> 0.2817 + 0.9767 * x_bar
[1] 4.494533
> # centered model: yij = 4.494533 + 0.9767(xj - x_bar) + error
> beta = c(4.494533,0.9767)
```

Hence, the centered model is

$$y_{ij} = 4.494533 + 0.9767(x_j - \bar{x}) + \varepsilon_{ij}.$$

To draw the chart of β_0 , β_1 , σ^2 (MSE), we need to calculate $\widehat{\beta}_{i0}$, $\widehat{\beta}_{i1}$ and

MSE_i.

```
> Sxx = sum(centered_x^2)
>
> # Calculate beta hat and MSE
> beta_hat = matrix(0,6,2)
> MSE = numeric(6)
> for(i in 1:6){
+   a = 1+3*(i-1)
+   b = 3*i
+   Y = y[a:b]
+   beta_hat[i,1] = mean(Y)
+   beta_hat[i,2] = sum(Y*centered_x)/Sxx
+   MSE[i] = sum((Y-beta_hat[i,1]-beta_hat[i,2]*centered_x)^2)
+ }
```

```

> beta_hat
      [,1]      [,2]
[1,] 4.573333 0.9862215
[2,] 4.470000 0.9692984
[3,] 4.510000 0.9823952
[4,] 4.603333 1.0406046
[5,] 4.513333 0.9935296
[6,] 4.523333 0.9826744
> MSE
[1] 8.627677e-03 4.234966e-03
[3] 3.137103e-03 7.032217e-02
[5] 1.750634e-03 8.097610e-06

```

Since $ARL = 200$, $\alpha^{overall} = 0.005$, $\alpha = 1 - (1 - \alpha^{overall})^{1/3} = 0.00167$ and $z_{\alpha/2} = 3.143492$.

```

> (alpha = 1-(1-0.005)^(1/3))
[1] 0.001669452
> sigma = 0.06826
> (z = qnorm(1-alpha/2))
[1] 3.143492

```

In addition, $UCL_{\beta_0} = \beta_0 + z_{\alpha/2} \sqrt{\frac{\sigma^2}{n}} = 4.6184$, $LCL_{\beta_0} = \beta_0 - z_{\alpha/2} \sqrt{\frac{\sigma^2}{n}} = 4.3706$,

$$UCL_{\beta_1} = \beta_1 + z_{\alpha/2} \sqrt{\frac{\sigma^2}{S_{xx}}} = 1.013, \quad LCL_{\beta_1} = \beta_1 - z_{\alpha/2} \sqrt{\frac{\sigma^2}{S_{xx}}} = 0.9402,$$

$$UCL_{MSE} = 0.0520 \text{ and } LCL_{MSE} = 5.0997 \times 10^{-9}$$

```

> se_beta = sigma * sqrt(c(1/3,1/Sxx))
>
> # information for chart for beta0
> U0 = beta[1]+z*se_beta[1]
> L0 = beta[1]-z*se_beta[1]
> cat("\nUCL = ",U0,"\nLCL = ",L0,
+     "\ncenter = ",beta[1]," \n\n")

UCL = 4.618418
LCL = 4.370648
center = 4.494533

>
> # information for chart for beta1
> U1 = beta[2]+z*se_beta[2]
> L1 = beta[2]-z*se_beta[2]
> cat("\nUCL = ",U1,"\nLCL = ",L1,
+     "\ncenter = ",beta[2]," \n\n")

```

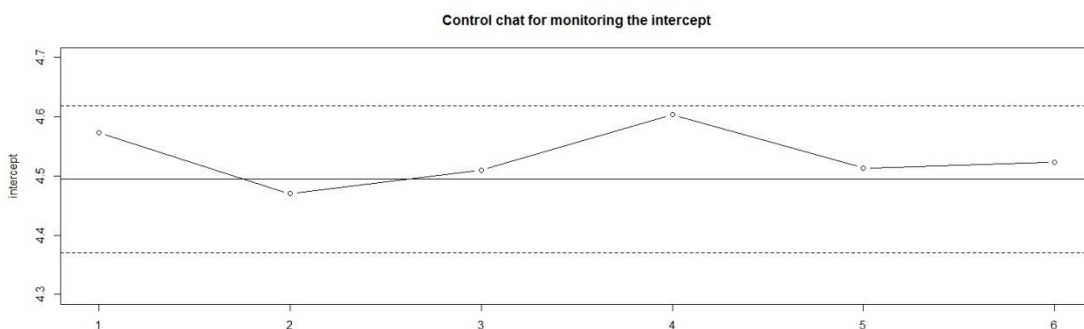
```
UCL = 1.013169
LCL = 0.9402314
center = 0.9767
```

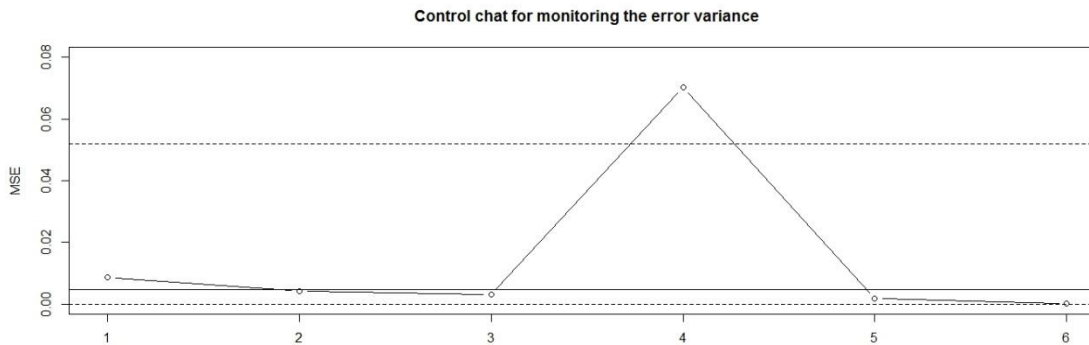
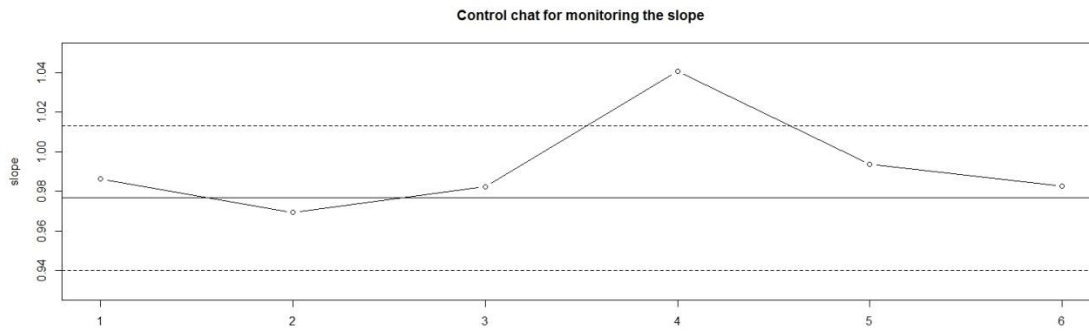
```
>
> # information for chart for sigma2
> U = sigma^2*qchisq(1-alpha/2,1)
> L = sigma^2*qchisq(alpha/2,1)
> cat("\nUCL = ",U,"\nLCL = ",L,
+     "\ncenter = ",sigma^2,"\n\n")
```

```
UCL = 0.05201032
LCL = 5.099652e-09
center = 0.004659428
```

Next, plot the chart of β_0 , β_1 , σ^2 (MSE).

```
> plot(beta_hat[,1], type = "b",
+ ylim = c(4.3,4.7),
+ xlab = "", ylab = "intercept",
+ main = "Control chat for monitoring the intervept")
> abline(U0,0,lty = 2)
> abline(L0,0,lty = 2)
> abline(beta[1],0)
> plot(beta_hat[,2], type = "b",
+ ylim = c(0.93,1.05),
+ xlab = "", ylab = "slope",
+ main = "Control chat for monitoring the slope")
> abline(U1,0,lty = 2)
> abline(L1,0,lty = 2)
> abline(beta[2],0)
> plot(MSE, type = "b",
+      ylim = c(0,0.025),
+      xlab = "", ylab = "MSE",
+      main = "Control chat for monitoring the error
variance")
> abline(U,0,lty = 2)
> abline(L,0,lty = 2)
> abline(sigma^2,0)
```





These figures are pretty similar to the figures in textbook, except the UCL_{MSE} (it's close to 0.005 in textbook).

2.

(1) Prove $\mathbf{X} \sim N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \rightarrow \mathbf{Z} = \boldsymbol{\Sigma}^{-1/2}(\mathbf{X} - \boldsymbol{\mu}) \sim N_p(0_p, I_p)$

Since $\boldsymbol{\Sigma}^{-1/2}(\mathbf{X} - \boldsymbol{\mu})$ is linear in \mathbf{X} , \mathbf{Z} is multivariate normal.

Moreover,

$$E[\mathbf{Z}] = E[\boldsymbol{\Sigma}^{-1/2}(\mathbf{X} - \boldsymbol{\mu})] = \boldsymbol{\Sigma}^{-1/2}(E[\mathbf{X}] - \boldsymbol{\mu}) = 0_p$$

$$\begin{aligned} \text{Cov}[\boldsymbol{\Sigma}^{-1/2}(\mathbf{X} - \boldsymbol{\mu})] &= \boldsymbol{\Sigma}^{-1/2} \text{Cov}[\mathbf{X} - \boldsymbol{\mu}] \boldsymbol{\Sigma}^{-1/2'} = \boldsymbol{\Sigma}^{-1/2} \text{Cov}[\mathbf{X}] \boldsymbol{\Sigma}^{-1/2'} \\ &= \boldsymbol{\Sigma}^{-1/2} \boldsymbol{\Sigma} \boldsymbol{\Sigma}^{-1/2'} = I_p \end{aligned}$$

Hence, $\mathbf{Z} \sim N_p(0_p, I_p)$.

(2) Prove $\mathbf{X}_1, \dots, \mathbf{X}_n \sim N_p(\boldsymbol{\mu}, \Sigma) \rightarrow \bar{\mathbf{X}} \sim N_p(\boldsymbol{\mu}, \frac{\Sigma}{n})$

Since $\bar{\mathbf{X}}$ is linear in $\mathbf{X}_1, \dots, \mathbf{X}_n$, $\bar{\mathbf{X}}$ is multivariate normal.

Moreover,

$$E[\bar{\mathbf{X}}] = \frac{1}{n} E[\mathbf{X}_1 + \dots + \mathbf{X}_n] = \frac{1}{n} \sum_{i=1}^n E[\mathbf{X}_i] = \frac{1}{n} \sum_{i=1}^n \boldsymbol{\mu} = \boldsymbol{\mu}$$

$$\begin{aligned} \text{Cov}[\bar{\mathbf{X}}] &= \frac{1}{n^2} \text{Cov}[\mathbf{X}_1 + \dots + \mathbf{X}_n] = \frac{1}{n^2} \sum_{i=1}^n \text{Cov}[\mathbf{X}_i] = \frac{1}{n^2} \sum_{i=1}^n \Sigma \\ &= \frac{\Sigma}{n} \end{aligned}$$

Hence, $\bar{\mathbf{X}} \sim N_p(\boldsymbol{\mu}, \frac{\Sigma}{n})$.

3.

(1)

Since $\Sigma = (1 - 0.5) I_3 + 0.5 \mathbf{1}\mathbf{1}'$,

$$\begin{aligned} \Sigma^{-1} &= \frac{1}{1-0.5} I_3 - \frac{0.5}{(1-0.5)[1+0.5(3-1)]} \mathbf{1}\mathbf{1}' = 2 I_3 - 0.5 \mathbf{1}\mathbf{1}' \\ &= \begin{pmatrix} 1.5 & -0.5 & -0.5 \\ -0.5 & 1.5 & -0.5 \\ -0.5 & -0.5 & 1.5 \end{pmatrix} \end{aligned}$$

(2)

```
library(MASS)
mu = c(-1,-1,-1)
sigma = 0.5*diag(1,3) + 0.5*rep(1,3)%*%t(rep(1,3))

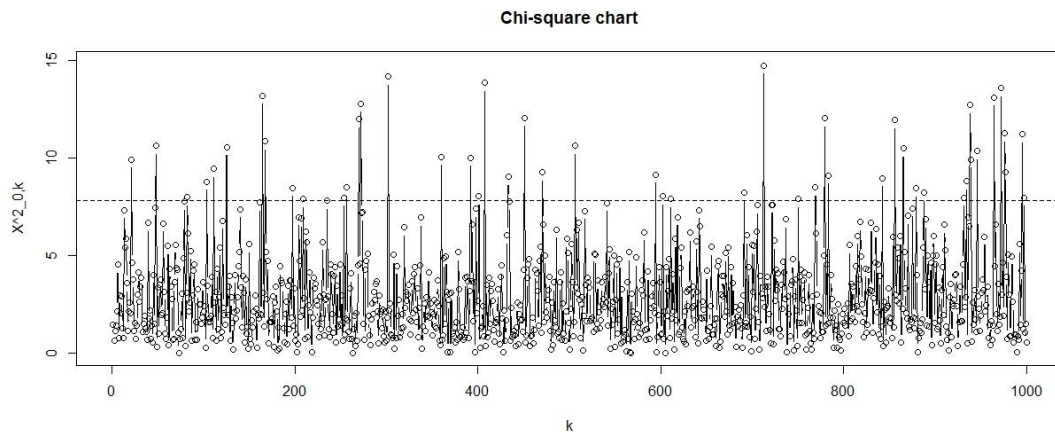
> set.seed(465)
chi = numeric(1000)
for(i in 1:1000){
  X = mvrnorm(25,mu,sigma)
  X_bar = colMeans(X)
  temp = X_bar - mu
  chi[i] = 25*t(temp)%*%solve(sigma)%*%(temp)
}
```

```

UCL = qchisq(0.95,3)

U = max(max(chi),UCL)
plot(chi , type = "b" , ylim = c(0,U+0.1) ,
      main = "Chi-square chart" , xlab = "k" ,
      ylab = "X^2_0,k")
abline(UCL , 0 , lty = 2)
> (ratio = sum(chi > UCL)/1000)
[1] 0.049

```



The ratio of out-of-control points is 0.049 and it's close to the nominal $\alpha=0.05$.

(3)

```

> library(MASS)
> mu = c(-1,-1,-1)
> sigma = 0.5*diag(1,3) + 0.5*rep(1,3)%*%t(rep(1,3))
>
> set.seed(465)
>
> X_bar_k = matrix(0,3,1000)
> Sk = array(0,dim=c(3,3,1000))
> for(i in 1:1000){
+   X = mvrnorm(25,mu,sigma)
+   X_bar_k[,i] = colMeans(X)
+   Sk[,,i] = cov(X)
+ }
>
> (X_bar_bar = rowMeans(X_bar_k))
[1] -1.013276 -1.010876 -1.005413
>
> (S = rowMeans(Sk , dim = 2))

```

	[,1]	[,2]	[,3]
[1,]	1.0244434	0.5029764	0.5072474
[2,]	0.5029764	1.0011144	0.5024538
[3,]	0.5072474	0.5024538	1.0066791

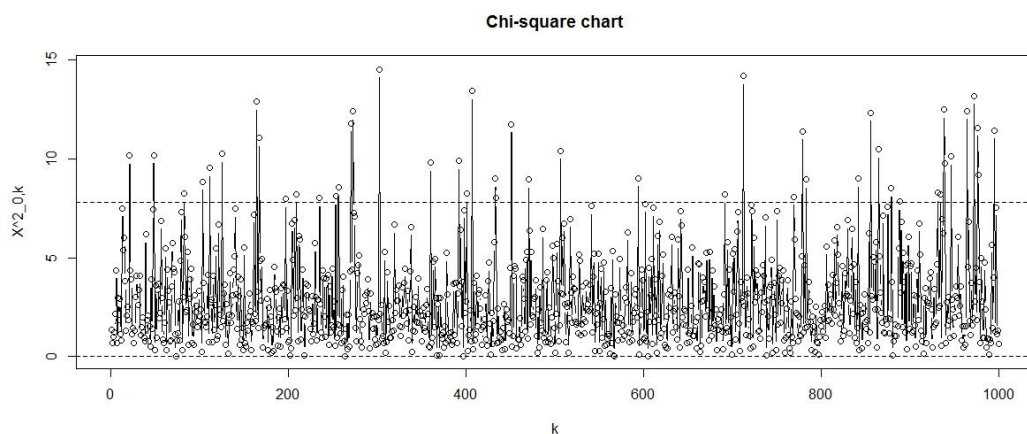
The $\bar{\bar{X}}$ and S seem to be close to the μ and Σ .

(4)

```

> chisq = numeric(1000)
> for(j in 1:1000){
+   temp = X_bar_k[j] - X_bar_bar
+   chisq[j] = 25*t(temp)%*%solve(S)%*%temp
+ }
>
> UCL = qchisq(0.95,3)
>
> U = max(max(chisq),UCL)
> plot(chisq , type = "b" , ylim = c(0,U+0.1) ,
+      main = "Chi-square chart" , xlab = "k" ,
+      ylab = "X^2_0,k")
> abline(UCL , 0 , lty = 2)
> abline(0 , 0 , lty = 2)
>
> (ratio = sum(chisq > UCL)/1000)
[1] 0.046

```



The ratio of out-of-control points is 0.046. It's close to the nominal $\alpha=0.05$, but slightly less than the value in (2).