

**HW #2 Quality Control, Spring 2013, [+10 points]**

**Due 4 / 9 (Tue) [ Submit after the Tuesday seminar, or put in my mailbox ]**

**Q1. Range method [+3]**

Let  $X_1, X_2 \stackrel{iid}{\sim} N(\mu, \sigma^2)$  and  $W = \frac{X_{(2)} - X_{(1)}}{\sigma}$  be the relative range.

- 1) Find the p.d.f. of  $W$ .
- 2) Calculate  $d_2 = E[W]$  and compare it with Table's value.
- 3) Do you know the name of this distribution of  $W$  ?  
(answer is not unique; clearly state why do you call its name)

**Q2. ARL [+2]**

The average run length is **numerically** obtained by the Monte Carlo method as follows:

$$E[L] \approx \frac{1}{M} \sum_{i=1}^M L_i, \quad M = 30000,$$

where  $L_1, \dots, L_M$  are iid replications of  $L$ .

- 1) Calculate the in-control ARL under the 3-sigma limits by Monte Carlo simulation (with codes). Compare your result with the **theoretical** value.
- 2) Do the same thing as 1) for  $SD[L]$ .

**Q3. P-chart [+2]** A manufacture conducts a waterproof testing for 5 electric boards. The number of defective circuits on the board is recorded as follows:

Board ID	1	2	3	4	5
The number of defectives circuits	20	20	23	25	12
The number of circuits	200	200	200	200	200

1. Draw a p-control chart.
2. The p-chart based on 3-sigma limits relies on the normal approximation. Please discuss whether the normal approximation to this dataset is suitable or not.  
(answer is not unique; give numerical or theoretical evidence for your answer)

**Q4. OC curve for P-chart [+2]**

Draw OC curve for  $n=50$  with  $LCL=0.0303$ ,  $UCL=0.3697$  by directly computing the Binomial probability masses (using computer). Compare your results with the OC curve based on hand writing.

## Quality Control QW #2

+10/10

Q1. +3

1) We have  $X_1, X_2 \stackrel{iid}{\sim} N(\mu, \sigma^2)$

Then  $\frac{X_1 - \mu}{\sigma}, \frac{X_2 - \mu}{\sigma} \stackrel{iid}{\sim} N(0, 1)$

And  $\max\{\frac{X_1 - \mu}{\sigma}, \frac{X_2 - \mu}{\sigma}\} = \frac{X_{(2)} - \mu}{\sigma}$ ,  $\min\{\frac{X_1 - \mu}{\sigma}, \frac{X_2 - \mu}{\sigma}\} = \frac{X_{(1)} - \mu}{\sigma}$

Let  $U = \frac{X_{(1)} - \mu}{\sigma}$  and  $V = \frac{X_{(2)} - \mu}{\sigma}$

→ The joint pdf of U, V is

$$f_{U,V}(u,v) = 2! [\phi(x)]^2 = 2 \frac{1}{2\pi} e^{-\frac{u^2+v^2}{2}} = \frac{1}{\pi} e^{-\frac{u^2+v^2}{2}}, -\infty < u < v < \infty \quad \checkmark$$

Again, let  $Y = U$  and  $W = V - U$

⇔  $U = Y, V = Y + W$  → The Jacobian is  $J = \begin{vmatrix} \frac{\partial U}{\partial Y} & \frac{\partial U}{\partial W} \\ \frac{\partial V}{\partial Y} & \frac{\partial V}{\partial W} \end{vmatrix} = \begin{vmatrix} 1 & 0 \\ 1 & 1 \end{vmatrix} = 1$

→ the joint pdf of Y, W is

$$\begin{aligned} f_{Y,W}(y,w) &= f_{U,V}(y, y+w) |1| = \frac{1}{\pi} e^{-\frac{y^2 + (y+w)^2}{2}} = \frac{1}{\pi} e^{-\frac{1}{2}(y^2 + y^2 + 2yw + w^2)} \\ &= \frac{1}{\pi} e^{-\frac{1}{2}w^2} e^{-(y^2 + wy)} = \frac{1}{\pi} e^{-\left(\frac{1}{2}w^2 - \left(\frac{w}{2}\right)^2\right)} e^{-(y^2 + wy + \left(\frac{w}{2}\right)^2)} \\ &= \frac{1}{\pi} e^{-\frac{w^2}{4}} e^{-(y + \frac{w}{2})^2}, -\infty < y < \infty, w > 0 \end{aligned}$$

So the marginal pdf of W is

$$\begin{aligned} f_W(w) &= \int_{-\infty}^{\infty} f_{Y,W}(y,w) dy = \int_{-\infty}^{\infty} \frac{1}{\pi} e^{-\frac{w^2}{4}} e^{-(y + \frac{w}{2})^2} dy = \frac{1}{\pi} e^{-\frac{w^2}{4}} \int_{-\infty}^{\infty} e^{-(y + \frac{w}{2})^2} dy \\ &= \frac{1}{\pi} e^{-\frac{w^2}{4}} \int_{-\infty}^{\infty} e^{-(y + \frac{w}{2})^2} d(y + \frac{w}{2}) = \frac{1}{\pi} e^{-\frac{w^2}{4}} \sqrt{\pi} \\ &= \frac{1}{\sqrt{\pi}} e^{-\frac{w^2}{4}}, w > 0 \end{aligned}$$

Let  $I = \int_{-\infty}^{\infty} e^{-x^2} dx$ .

$$\begin{aligned} \Rightarrow I^2 &= \int_{-\infty}^{\infty} e^{-x^2} dx \int_{-\infty}^{\infty} e^{-y^2} dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)} dx dy \\ &= \int_0^{\infty} \int_0^{2\pi} e^{-r^2} (r dr d\theta) = \int_0^{\infty} r e^{-r^2} \int_0^{2\pi} d\theta dr \\ &= 2\pi \int_0^{\infty} r e^{-r^2} dr = 2\pi \left(-\frac{1}{2} e^{-r^2} \Big|_0^{\infty}\right) \\ &= \pi \\ \Rightarrow I &= \sqrt{\pi} \end{aligned}$$



2)

$$d_2 = E[W] = \int_0^{\infty} w \frac{1}{\sqrt{\pi}} e^{-\frac{w^2}{4}} dw = \frac{1}{\sqrt{\pi}} \int_0^{\infty} w e^{-\frac{w^2}{4}} dw = \frac{1}{\sqrt{\pi}} (2e^{-\frac{w^2}{4}} \Big|_0^{\infty}) = \frac{1}{\sqrt{\pi}} 2 \doteq 1.128$$

The value of  $d_2$  table at  $n=2$  is 1.128

So we find that the result is similar to the table's value. ✓

3)

$$\text{We have } f_W(w) = \frac{1}{\sqrt{\pi}} e^{-\frac{w^2}{4}}, w > 0$$

$$\text{And we can rewrite it as } f_W(w) = 2 \frac{1}{2\sqrt{\pi}} e^{-\frac{w^2}{2 \times 2}}, w > 0$$

It's similar to the pdf of  $N(0,2)$ .

But  $W$  is with positive support and its pdf is twice as much as the pdf of  $N(0,2)$ .

So I call it as "positive normal". (Because I think "twice" is used to adjust the pdf such that its integration in  $(0, \infty)$  equal to 1) ✓

Q2. +2/2

1) We have  $X_{ij}, i=1, \dots, m, j=1, \dots, n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

$$\Rightarrow \bar{X}_i, i=1, \dots, m \stackrel{iid}{\sim} N\left(\mu, \frac{\sigma^2}{n}\right)$$

$$\Rightarrow \frac{\bar{X}_i - \mu}{\sigma/\sqrt{n}} \stackrel{iid}{\sim} N(0, 1) \quad \checkmark$$

And we say that process is out-of-control if  $\bar{X}_i > \mu_0 + 3\frac{\sigma}{\sqrt{n}}$  or  $\bar{X}_i < \mu_0 - 3\frac{\sigma}{\sqrt{n}}$

$$\Leftrightarrow \frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} > 3 \text{ or } \frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} < -3$$

So generate some random variates from  $N(0, 1)$  to simulate in-control ARL

```
> run_length=function(n)#use n-sigma
+ {
+ k=1
+ repeat{
+ x=rnorm(1)
+ if(abs(x)>n)break
+ k=k+1
+ }
+ return(k)
+ }
>
> L=numeric()#generate Li
> for(i in 1:30000){
+ temp=run_length(3)
+ L[i]=temp
+ }
> mean(L)#average Li
[1] 372.1721
```

And the theoretical value of in-control ARL is

$$\begin{aligned} E[L | \mu = \mu_0] &= \frac{1}{P(\bar{X}_i > \mu_0 + 3\frac{\sigma}{\sqrt{n}} \text{ or } \bar{X}_i < \mu_0 - 3\frac{\sigma}{\sqrt{n}} | \mu = \mu_0)} = \frac{1}{P\left(\frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} > 3 \text{ or } \frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} < -3 | \mu = \mu_0\right)} \\ &= \frac{1}{1 - \Phi(3) + \Phi(-3)} \doteq \frac{1}{0.0027} \doteq 370 \end{aligned}$$

So the two values are close. ✓



2)

We use the same data from 1)

And then calculate the standard deviation of the data

> sqrt(var(L))#SD of Li's

[1] 371.5853

Also, the theoretical value of  $SD_0[L]$  is

$$SD_0[L] = \frac{\sqrt{1-p}}{p}$$

where  $p = P\left(\frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} > 3 \text{ or } \frac{\bar{X}_i - \mu_0}{\sigma/\sqrt{n}} < -3 \mid \mu = \mu_0\right) = 1 - \Phi(3) + \Phi(-3) = 0.0027$

$\Rightarrow$  the theoretical value of  $SD_0[L]$  is  $\frac{\sqrt{1-0.0027}}{0.0027} \doteq 369.87$

We find that two values are close.

Q3. +2/2

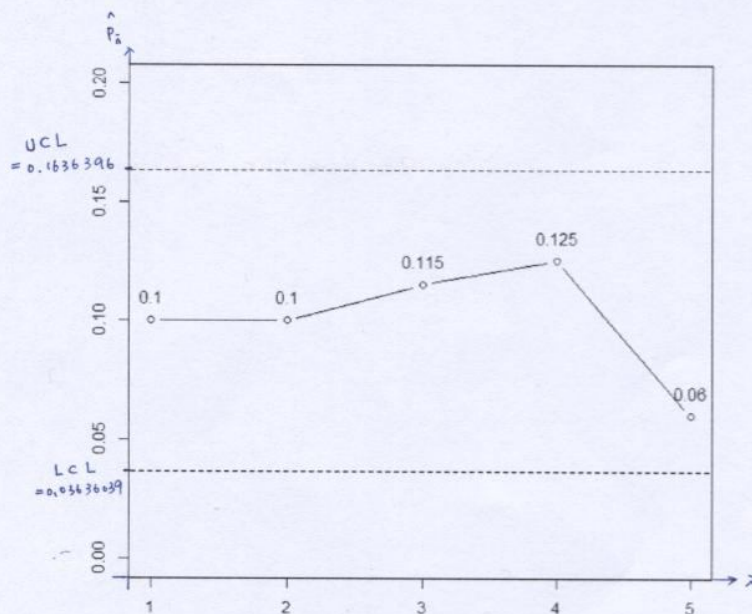
1.

n=200	$D_i$	$\hat{p}_i$
1	20	0.1
2	20	0.1
3	23	0.115
4	25	0.125
5	12	0.06

$\Rightarrow \bar{p}=0.1$

$$UCL = \bar{p} + 3 \sqrt{\frac{\bar{p}(1-\bar{p})}{n}} = 0.1 + 3 \sqrt{\frac{0.1(1-0.1)}{200}} = 0.1636396 \quad \checkmark$$

$$LCL = \bar{p} - 3 \sqrt{\frac{\bar{p}(1-\bar{p})}{n}} = 0.1 - 3 \sqrt{\frac{0.1(1-0.1)}{200}} = 0.03636039 \quad \checkmark$$



2. Reference?  
In the course of biostatistics, many theorem or lemma of binomial distribution which use normal approximation are required  $n \cdot p \cdot (1-p) \geq 5$ .  
So I calculate  $n \cdot p \cdot (1-p)$  in this case and the result is  $200 \times 0.1 \times 0.9 = 18$ .  
18 is larger than 5.  
So I think the normal approximation is suitable in this case.  $\checkmark$



Q4. +2/2

We have  $\beta = P(LCL \leq \hat{p}_i \leq UCL | H_1: p \neq p_0) = P(n \times LCL \leq D_i \leq n \times UCL | H_1: p \neq p_0)$

$$= \sum_{x=[n \times LCL]+1}^{[n \times UCL]} \binom{n}{x} p^x (1-p)^{n-x}$$

Here,  $n=50$

$$[n \times UCL] = [50 \times 0.3697] = [18.485] = 18$$

$$[n \times LCL] = [50 \times 0.0303] = [1.515] = 1$$

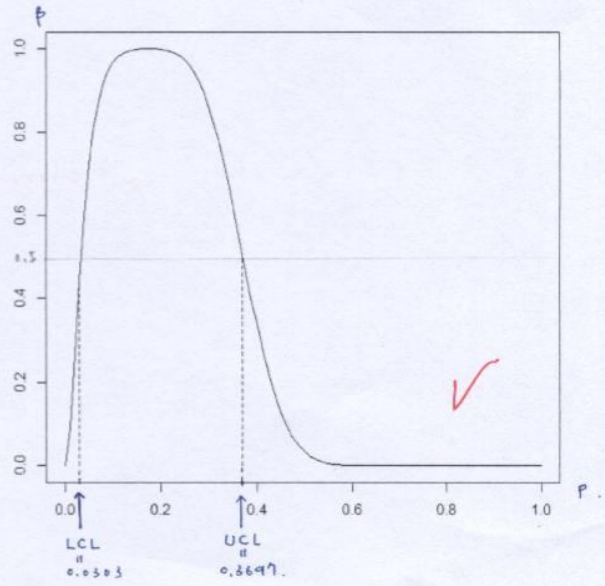
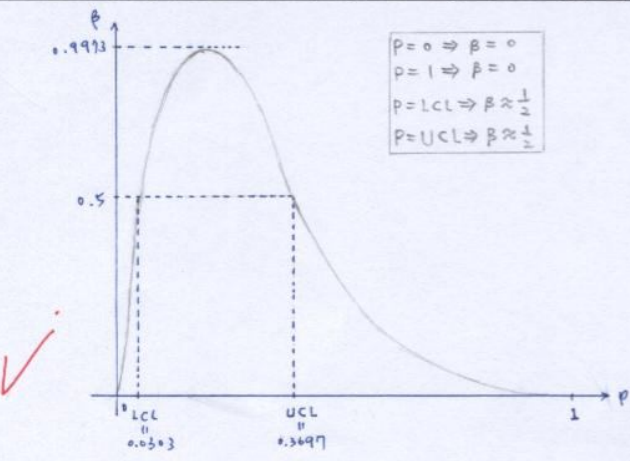
So make a code as following.

```
> beta=function(p)                                     #a function to calculate beta w.r.t p
+ {
+ lfn=log(factorial(50))
+ lp=log(p)
+ l1_p=log(1-p)
+ temp=numeric()
+ for(i in 2:18){                                     #the log value of pdf of b(50,p) with x=i
+ lfn_x=log(factorial(50-i))
+ lfx=log(factorial(i))
+ temp[i-1]=lfn-lfn_x-lfx+i*lp+(n-i)*l1_p
+ }
+ temp=exp(temp)
+ return(sum(temp))
+ }
```

why do you use "log" ?

And then draw the OC curve for  $n=50$  by using computer

```
> x=seq(0,1,by=0.01)
> y=numeric()
> for(i in 1:101){
+ y[i]=beta(x[i])
+ }
> plot(x,y,xlim=c(0,1),ylim=c(0,1),type="l")
> lines(rep(0.0303,2),c(0,beta(0.0303)),lty=2)
> lines(rep(0.3697,2),c(0,beta(0.3697)),lty=2)
```



In hand-writing one, we draw the line only by 4 points,  $(0,0)$ ,  $(1,1)$ ,  $(0.0303, \frac{1}{2})$ ,  $(0.3697, \frac{1}{2})$ . But in the other one, it's drawn by 101 points. So the OC curve drawn by computer is more precise.



1) Q3: +3/2  
<Code>

```
> library(qcc)
Loading required package: MASS
Package 'qcc', version 2.2
Type 'citation("qcc")' for citing this R package in publications.
> circuits=c(20,20,23,25,12)
> circuits.p=qcc(data=circuits,sizes=200,type="p")
> summary(circuits.p)
```

Call:  
qcc(data = circuits, type = "p", sizes = 200)

p chart for circuits

Summary of group statistics:

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
0.060	0.100	0.100	0.100	0.115	0.125

Group sample size: 200  
Number of groups: 5  
Center of group statistics: 0.1  
Standard deviation: 0.3

(+1)

Control limits:  
LCL            UCL  
0.03636039 0.1636396

### p-control chart

