

Homework#1

Name: 范又方 Fan, You-fang

Student number: 106225022

Exercise 6.3 [P300]

Let X_1, \dots, X_n be a random sample from the pdf

$$f(x|\mu, \sigma) = \frac{1}{\sigma} e^{-\frac{x-\mu}{\sigma}}, \quad \mu < x < \infty, 0 < \sigma < \infty.$$

Find a two-dimensional sufficient statistic for (μ, σ) .

Sol :

The joint pdf of the sample \mathbf{X} is

$$f(\mathbf{x}|\mu, \sigma) = \prod_{i=1}^n \frac{1}{\sigma} \exp\left(-\frac{x_i - \mu}{\sigma}\right), \quad \mu < x_i < \infty, 0 < \sigma < \infty, i = 1, 2, \dots, n$$

We can use the indicator function. $I_A(x)$ is the indicator function of the set A ; that is, it is equal to 1 if $x \in A$ and equal to 0 otherwise. Now $\mu < x_i < \infty$, $i = 1, 2, \dots, n$ if and only if $\min\{x_1, x_2, \dots, x_n\} > \mu$.

So we have

$$\begin{aligned} f(\mathbf{x}|\mu, \sigma) &= \prod_{i=1}^n \frac{1}{\sigma} \exp\left(-\frac{x_i - \mu}{\sigma}\right) I_{(\mu, \infty)}(x_i) \\ &= \frac{1}{\sigma^n} \exp\left(-\frac{\sum_{i=1}^n (x_i - \mu)}{\sigma}\right) I_{(-\infty, X_{(1)})}(\mu) \\ &= \frac{1}{\sigma^n} \exp\left(\frac{n\mu}{\sigma}\right) \exp\left(-\frac{\sum_{i=1}^n x_i}{\sigma}\right) I_{(-\infty, X_{(1)})}(\mu) \end{aligned}$$

Since $f(\mathbf{x}|\mu, \sigma) = h(\mathbf{x})g(T(\mathbf{X})|\mu, \sigma)$, where $h(\mathbf{x}) = 1$ and

$$g(T(\mathbf{X})|\mu, \sigma) = \frac{1}{\sigma^n} \exp\left(\frac{n\mu}{\sigma}\right) \exp\left(-\frac{\sum_{i=1}^n x_i}{\sigma}\right) I_{(-\infty, X_{(1)})}(\mu), \quad T_1(\mathbf{X}) = X_{(1)}, \quad T_2(\mathbf{X}) =$$

$$\sum_{i=1}^n X_i$$

Therefore, by the Factorization Theorem, $(X_{(1)}, \sum_{i=1}^n X_i)$ is a two-dimensional sufficient statistic for (μ, σ) .

Exercise 6.5 [P300]

Let X_1, \dots, X_n be independent random variables with pdfs

$$f(x_i|\theta) = \begin{cases} \frac{1}{2i\theta} & -i(\theta - 1) < x_i < i(\theta + 1) \\ 0 & \text{otherwise,} \end{cases}$$

where $\theta > 0$. Find a two-dimensional sufficient statistic for θ .

Sol :

The joint pdf of the sample \mathbf{X} is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^n \frac{1}{2i\theta}, \quad -i(\theta - 1) < x_i < i(\theta + 1), \quad \theta > 0$$

Now $-i(\theta - 1) < x_i < i(\theta + 1)$, using the indicator function, we can see the inequality is $-i(\theta - 1) < \min x_i$ and $\max x_i < i(\theta + 1)$.

$$\Rightarrow \begin{cases} -i(\theta - 1) < \min x_i \\ i(\theta + 1) > \max x_i \end{cases} = \begin{cases} \theta - 1 > -\min \frac{x_i}{i} \\ \theta + 1 > \max \frac{x_i}{i} \end{cases} = \begin{cases} \theta > 1 - \min \frac{x_i}{i} \\ \theta > \max \frac{x_i}{i} - 1 \end{cases}$$

So we have

$$\begin{aligned} f(\mathbf{x}|\theta) &= \prod_{i=1}^n \frac{1}{2i\theta} I_{(-i(\theta-1), i(\theta+1))}(x_i) \\ &= \left(\frac{1}{2\theta}\right)^n \prod_{i=1}^n \frac{1}{i} I_{(1-\min \frac{x_i}{i}, \infty)}(\theta) I_{(\max \frac{x_i}{i}-1, \infty)}(\theta) \end{aligned}$$

Since $f(x|\mu, \sigma) = h(x)g(T(\mathbf{X})|\mu, \sigma)$, where $h(x) = \prod_{i=1}^n \frac{1}{i}$ and $g(T(\mathbf{X})|\theta) =$

$$\left(\frac{1}{2\theta}\right)^n I_{(1-\min \frac{x_i}{i}, \infty)}(\theta) I_{(\max \frac{x_i}{i}-1, \infty)}(\theta), \quad T_1(\mathbf{X}) = \min \frac{x_i}{i}, \quad T_2(\mathbf{X}) = \max \frac{x_i}{i}$$

Therefore, by the Factorization Theorem, $(\min \frac{x_i}{i}, \max \frac{x_i}{i})$ is a two-dimensional sufficient statistic for θ .

Exercise 6.6 [P300]

Let X_1, \dots, X_n be a random sample from a gamma(α, β) population. Find a two-dimensional sufficient statistic for (α, β) .

Sol :

The joint pdf of the sample \mathbf{X} is

$$\begin{aligned} f(\mathbf{x}|\alpha, \beta) &= \prod_{i=1}^n \frac{1}{\Gamma(\alpha)\beta^\alpha} x_i^{\alpha-1} \exp\left(-\frac{x_i}{\beta}\right) \\ &= \left(\frac{1}{\Gamma(\alpha)\beta^\alpha}\right)^n (\prod_{i=1}^n x_i)^{\alpha-1} \exp\left(-\frac{\sum_{i=1}^n x_i}{\beta}\right) \end{aligned}$$

Since $f(x|\alpha, \beta) = h(x)g(T(\mathbf{X})|\alpha, \beta)$, where $h(x) = 1$ and $g(T(\mathbf{X})|\alpha, \beta) =$

$$\left(\frac{1}{\Gamma(\alpha)\beta^\alpha}\right)^n (\prod_{i=1}^n x_i)^{\alpha-1} \exp\left(-\frac{\sum_{i=1}^n x_i}{\beta}\right) \quad , \quad T_1(\mathbf{X}) = \prod_{i=1}^n X_i \quad , \quad T_2(\mathbf{X}) = \sum_{i=1}^n X_i$$

Therefore, by the Factorization Theorem, $(\prod_{i=1}^n X_i, \sum_{i=1}^n X_i)$ is a two-dimensional sufficient statistic for (α, β) .

Exercise 6.7 [P300]

Let $f(x, y|\theta_1, \theta_2, \theta_3, \theta_4)$ be the bivariate pdf for the uniform distribution on the rectangle with lower left corner (θ_1, θ_2) and upper right corner (θ_3, θ_4) in \mathfrak{R}^2 . The parameters satisfy $\theta_1 < \theta_3$ and $\theta_2 < \theta_4$. Let $(X_1, Y_1), \dots, (X_n, Y_n)$ be a random sample from this pdf. Find a four-dimensional sufficient statistic for $(\theta_1, \theta_2, \theta_3, \theta_4)$.

Sol :

The joint pdf is

$$f(\mathbf{x}, \mathbf{y}|\theta) = \prod_{i=1}^n \frac{1}{\theta_3 - \theta_1} \frac{1}{\theta_4 - \theta_2} \quad , \quad \theta_1 \leq x_i \leq \theta_3 \quad , \quad \theta_2 \leq y_i \leq \theta_4$$

Using the indicator function,

$$\begin{aligned} \theta_1 &\leq \min\{x_1, \dots, x_n\} = x_{(1)} \quad , \quad \max\{x_1, \dots, x_n\} = x_{(n)} \leq \theta_3 \quad , \\ \theta_2 &\leq \min\{y_1, \dots, y_n\} = y_{(1)} \quad , \quad \max\{y_1, \dots, y_n\} = y_{(n)} \leq \theta_4 \end{aligned}$$

So we have

$$\begin{aligned} f(\mathbf{x}, \mathbf{y}|\theta) &= \prod_{i=1}^n \frac{1}{\theta_3 - \theta_1} \frac{1}{\theta_4 - \theta_2} I_{(\theta_1, \theta_3)}(x_i) I_{(\theta_2, \theta_4)}(y_i) \\ &= \left(\frac{1}{\theta_3 - \theta_1}\right)^n \left(\frac{1}{\theta_4 - \theta_2}\right)^n I_{(-\infty, x_{(1)})}(\theta_1) I_{(-\infty, y_{(1)})}(\theta_2) I_{(x_{(n)}, \infty)}(\theta_3) I_{(y_{(n)}, \infty)}(\theta_4) \end{aligned}$$

Since $f(x, y|\theta) = h(x)g(T(\mathbf{X})|\theta)$, where $h(x) = 1$ and

$g(T(\mathbf{X})|\theta) =$

$$\left(\frac{1}{\theta_3 - \theta_1}\right)^n \left(\frac{1}{\theta_4 - \theta_2}\right)^n I_{(-\infty, x_{(1)})}(\theta_1) I_{(-\infty, y_{(1)})}(\theta_2) I_{(x_{(n)}, \infty)}(\theta_3) I_{(y_{(n)}, \infty)}(\theta_4) \quad ,$$

$$T_1(\mathbf{X}) = X_{(1)} \quad , \quad T_2(\mathbf{X}) = Y_{(1)} \quad , \quad T_3(\mathbf{X}) = X_{(n)} \quad , \quad T_4(\mathbf{X}) = Y_{(n)}.$$

Therefore, by the Factorization Theorem, $(X_{(1)} , Y_{(1)} , X_{(n)} , Y_{(n)})$ is a four-dimensional sufficient statistic for $(\theta_1, \theta_2, \theta_3, \theta_4)$.

Exercise 6.8 [P301]

Let X_1, \dots, X_n be a random sample from a population with location pdf $f(x - \theta)$. Show that the order statistics, $T(X_1, \dots, X_n) = (X_{(1)}, \dots, X_{(n)})$, are a sufficient statistic for θ and no further reduction is possible.

Sol :

We know $X_1, \dots, X_n \sim f(x - \theta)$, $\theta \in \mathbb{R}$.

The joint pdf is

$$f(\mathbf{x}|\theta) = \prod_{i=1}^n f(x_i|\theta) = \prod_{i=1}^n f(x_i - \theta) = \prod_{i=1}^n f(x_{(i)} - \theta)$$

Where $x_{(1)} \leq x_{(2)} \leq \dots \leq x_{(n)}$ are the order statistics of X_1, \dots, X_n .

Since $f(x, y|\theta) = h(x)g(T(\mathbf{X})|\theta)$, where $h(x) = 1$ and $g(T(\mathbf{X})|\theta) = \prod_{i=1}^n f(x_{(i)} - \theta)$, $T_i(\mathbf{X}) = X_{(i)}$, $i = 1, \dots, n$

Therefore, by the Factorization Theorem, $(X_{(1)}, \dots, X_{(n)})$ are a sufficient statistic for θ .

That no further reduction is possible without further restrictions on f . In this case, it suffices to notice the ratio

$$\frac{\prod_{i=1}^n f(x_{(i)} - \theta)}{\prod_{i=1}^n f(y_{(i)} - \theta)}$$

is general independent of θ only when $T(\mathbf{X}) = T(\mathbf{Y})$.