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Midterm exam2, High-dimensional Data Analysis, 2018 Spring [+32 points]

Not only answer but also calculation

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Q1 [+6] Consider a model

$$Y_i = \beta_0 + \sum_{j=1}^p \beta_j x_{ij} + \varepsilon_i, \quad \varepsilon_i \sim N(0, \sigma^2), \quad i=1, \dots, n.$$

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(1) [+3] Derive the Lasso estimator ($\hat{\beta}_\lambda, \lambda=?$) as a posterior mode.

(define a prior density and derive a posterior density)

Let $\beta_j \sim$ double exponential (b) $\forall j=1, 2, \dots, p \Rightarrow f(\beta_j) = \frac{1}{2b} \exp(-\frac{|\beta_j|}{b}) \quad \forall j=1, 2, \dots, p$

And because of $Y_i | \beta \sim N(\beta_0 + \sum_{j=1}^p \beta_j x_{ij}, \sigma^2) \quad \forall i=1, 2, \dots, n.$

$$\begin{aligned} \Rightarrow f(\beta | X) &\propto f(X | \beta) f(\beta) \\ &\propto \left[\prod_{i=1}^n \exp\left(-\frac{1}{2\sigma^2} (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2\right) \right] \times \left[\prod_{j=1}^p \exp\left(-\frac{|\beta_j|}{b}\right) \right] \\ &= \exp\left[\frac{1}{2\sigma^2} (y - X\beta)^T (y - X\beta) - \frac{1}{b} \sum_{j=1}^p |\beta_j| \right] \quad (\text{where } y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}, \beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}, X = \begin{bmatrix} x_{11} & \dots & x_{1p} \\ \vdots & & \vdots \\ x_{n1} & \dots & x_{np} \end{bmatrix}) \\ &= \exp\left[\frac{1}{2\sigma^2} (y - X\beta)^T (y - X\beta) + \frac{2\sigma^2}{b} \sum_{j=1}^p |\beta_j| \right] \\ \Rightarrow \hat{\beta}_\lambda &= \arg \max_{\beta} f(\beta | X) = \arg \min_{\beta} \left[(y - X\beta)^T (y - X\beta) + \frac{2\sigma^2}{b} \sum_{j=1}^p |\beta_j| \right] \end{aligned}$$

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(2) [+3] Derive the ridge estimator ($\hat{\beta}_\lambda, \lambda=?$) as a posterior mode.

\Rightarrow When $\lambda = \frac{2\sigma^2}{b}$, $\hat{\beta}_\lambda$ are equivalent to the Lasso estimator

(define a prior density and derive a posterior density)

Let $\beta_j \sim N(0, c) \quad \forall j=1, 2, \dots, p$

$\therefore Y_i | \beta \sim N(\beta_0 + \sum_{j=1}^p \beta_j x_{ij}, \sigma^2)$

$$\therefore f(\beta | X) \propto f(X | \beta) f(\beta) \propto \left[\prod_{i=1}^n \exp\left(-\frac{(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2}{2\sigma^2}\right) \right] \times \left[\prod_{j=1}^p \exp\left(-\frac{\beta_j^2}{2c}\right) \right]$$

$$= \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 - \frac{1}{2c} \sum_{j=1}^p \beta_j^2\right)$$

$$= \exp\left(-\frac{1}{2\sigma^2} (y - X\beta)^T (y - X\beta) - \frac{1}{2c} \sum_{j=1}^p \beta_j^2\right) \quad (\text{where } y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}, \beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}, X = \begin{bmatrix} x_{11} & \dots & x_{1p} \\ \vdots & & \vdots \\ x_{n1} & \dots & x_{np} \end{bmatrix})$$

$$= \exp\left(-\frac{1}{2\sigma^2} \left[(y - X\beta)^T (y - X\beta) + \frac{\sigma^2}{c} \sum_{j=1}^p \beta_j^2 \right]\right)$$

$$\Rightarrow \hat{\beta}_\lambda = \arg \max_{\beta} f(\beta | X) = \arg \max_{\beta} \exp\left(-\frac{1}{2\sigma^2} \left[(y - X\beta)^T (y - X\beta) + \frac{\sigma^2}{c} \sum_{j=1}^p \beta_j^2 \right]\right)$$

$$= \arg \min_{\beta} \left[(y - X\beta)^T (y - X\beta) + \frac{\sigma^2}{c} \sum_{j=1}^p \beta_j^2 \right]$$

\Rightarrow When $\lambda = \frac{\sigma^2}{c}$ (or $\lambda = \frac{\sigma^2}{c}$), this $\hat{\beta}_\lambda$ will be equivalent to the ridge estimator.

Moreover, $\hat{\beta}_\lambda$ can be represented as $(X^T X + \lambda I)^{-1} X^T y$.

+12 Q2 [+12] Consider a model without an intercept:

$$Y_i = \beta_1 x_{i1} + \beta_2 x_{i2} + \varepsilon_i, \quad \varepsilon_i \sim N(0, \sigma^2), \quad i = 1, \dots, n. \text{ Assume } \sum_{i=1}^n x_{i1} = \sum_{i=1}^n x_{i2} = 0 \text{ and } \sum_{i=1}^n x_{i1}^2 = \sum_{i=1}^n x_{i2}^2 = 1.$$

Answer the questions by using $r_{12} = \sum_{i=1}^n x_{i1}x_{i2}$, $r_{1y} = \sum_{i=1}^n x_{i1}Y_i$, and $r_{2y} = \sum_{i=1}^n x_{i2}Y_i$.

+1 (1) [+1] $X^T X = \begin{bmatrix} 1 & r_{12} \\ r_{12} & 1 \end{bmatrix}_{2 \times 2}$

$$X = \begin{bmatrix} x_{11} & x_{12} \\ \vdots & \vdots \\ x_{n1} & x_{n2} \end{bmatrix} \rightarrow X^T X = \begin{bmatrix} x_{11} & \dots & x_{n1} \\ x_{12} & \dots & x_{n2} \end{bmatrix} \begin{bmatrix} x_{11} & x_{12} \\ \vdots & \vdots \\ x_{n1} & x_{n2} \end{bmatrix} = \begin{bmatrix} \sum x_{i1}^2 & \sum x_{i1}x_{i2} \\ \sum x_{i1}x_{i2} & \sum x_{i2}^2 \end{bmatrix}$$

$$\Rightarrow (X^T X)^{-1} = \frac{1}{1-r_{12}^2} \begin{bmatrix} 1 & -r_{12} \\ -r_{12} & 1 \end{bmatrix}$$

$$\Rightarrow \text{LSE: } (X^T X)^{-1} X^T y = \frac{1}{1-r_{12}^2} \begin{bmatrix} 1 & -r_{12} \\ -r_{12} & 1 \end{bmatrix} \begin{bmatrix} r_{1y} \\ r_{2y} \end{bmatrix}$$

+2 (2) [+2] Derive the LSE

$$\hat{\beta}_1 = \frac{r_{1y} - r_{12} r_{2y}}{1 - r_{12}^2}$$

$$\hat{\beta}_2 = \frac{r_{2y} - r_{12} r_{1y}}{1 - r_{12}^2}$$

$$= \frac{1}{1-r_{12}^2} \begin{bmatrix} r_{1y} - r_{12} r_{2y} \\ r_{2y} - r_{12} r_{1y} \end{bmatrix}$$

Below, we assume $r_{12} = r_{1y} = r_{2y} = 1/2$. $\Rightarrow X^T X = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}$ and LSE: $\hat{\beta} = \begin{bmatrix} 1/3 \\ 1/3 \end{bmatrix}$.

+2 (3) [+2] Derive the ridge estimators $\hat{\beta}_{1\lambda}$ and $\hat{\beta}_{2\lambda}$

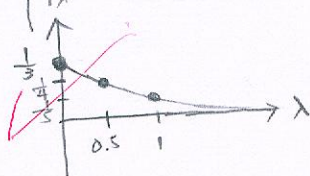
$$\hat{\beta}_\lambda = (X^T X + \lambda I)^{-1} X^T y = \begin{bmatrix} 1+\lambda & r_{12} \\ r_{12} & 1+\lambda \end{bmatrix}^{-1} \begin{bmatrix} r_{1y} \\ r_{2y} \end{bmatrix} \stackrel{r_{12}=r_{1y}=r_{2y}=1/2}{=} \begin{bmatrix} 1+\lambda & 0.5 \\ 0.5 & 1+\lambda \end{bmatrix}^{-1} \begin{bmatrix} 1/2 \\ 1/2 \end{bmatrix}$$

$$= \frac{4}{(2\lambda+1)(2\lambda+3)} \begin{bmatrix} 1+\lambda & -0.5 \\ -0.5 & 1+\lambda \end{bmatrix} \begin{bmatrix} 1 \\ 1 \end{bmatrix} \times \frac{1}{2} = \frac{2}{(2\lambda+1)(2\lambda+3)} \begin{bmatrix} \lambda+0.5 \\ \lambda+0.5 \end{bmatrix} = \frac{1}{2\lambda+3} \begin{bmatrix} 1 \\ 1 \end{bmatrix}$$

$$\Rightarrow \hat{\beta}_{1\lambda} = \hat{\beta}_{2\lambda} = \frac{1}{2\lambda+3} \text{ when } r_{12} = r_{1y} = r_{2y} = \frac{1}{2}$$

+2 (4) [+2] Draw the ridge trace for $\hat{\beta}_{1\lambda}$

λ	0	0.5	1	...	$\rightarrow \infty$
$\hat{\beta}_{1\lambda}$	$\frac{1}{3}$	$\frac{1}{4}$	$\frac{1}{5}$		0



this curve is decreasing and bounded at 0.

+2 (5) [+2] Derive the degrees of freedom df_λ

$$df_\lambda = \text{tr}[X(X^T X + \lambda I)^{-1} X^T] = \text{tr}[X^T X (X^T X + \lambda I)^{-1}] = \text{tr}\left[\begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix} \cdot \frac{4}{(2\lambda+1)(2\lambda+3)} \begin{bmatrix} 1+\lambda & -0.5 \\ -0.5 & 1+\lambda \end{bmatrix}\right]$$

$$= \frac{4}{(2\lambda+1)(2\lambda+3)} \text{tr}\begin{bmatrix} \lambda+0.75 & 0.5\lambda \\ 0.5\lambda & \lambda+0.75 \end{bmatrix} = \frac{4}{(2\lambda+1)(2\lambda+3)} \times (2\lambda+1.5) = \frac{8\lambda+6}{(2\lambda+1)(2\lambda+3)}$$

+1 (6) [+1] Derive the degrees of freedom df_λ when $\lambda=0$

$$df_{\lambda=0} = \frac{8\lambda+6}{(2\lambda+1)(2\lambda+3)} \Big|_{\lambda=0} = \frac{6}{1 \times 3} = 2$$

+1 (7) [+1] Derive the degrees of freedom df_λ when $\lambda \rightarrow \infty$

$$df_{\lambda \rightarrow \infty} = \lim_{\lambda \rightarrow \infty} \frac{8\lambda+6}{(2\lambda+1)(2\lambda+3)} = \lim_{\lambda \rightarrow \infty} \frac{8\lambda+6}{4\lambda^2+8\lambda+3} = \lim_{\lambda \rightarrow \infty} \frac{\frac{8}{\lambda} + \frac{6}{\lambda^2}}{4 + \frac{8}{\lambda} + \frac{3}{\lambda^2}} = \frac{0}{4} = 0$$

+1 (8) [+1] Choose λ by setting $df_\lambda = 1$ and $r_{12} = \frac{1}{2}$.

That is, solving that $\frac{8\lambda+6}{4\lambda^2+8\lambda+3} = 1 \Rightarrow 4\lambda^2+8\lambda+3 = 8\lambda+6$

$$\Rightarrow 4\lambda^2 = 3 \Rightarrow \lambda = \frac{\sqrt{3}}{2} \text{ (negative value is improper)}$$

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Q3 [+14] Let $f(x) = \sum_{j=1}^K \beta_j M_j(x)$, where $M_j(x)$'s are the spline basis functions. For a knot

sequence $\xi_1 < \xi_2 < \xi_3$ with $\Delta = \xi_2 - \xi_1 = \xi_3 - \xi_2$, let $z_i(t) = (t - \xi_i) / \Delta$ for $i = 1, 2$, and 3.

+2 (1) [+2] Show that $\int f''(x)^2 dx$ can be written as a quadratic form for $\beta = (\beta_1, \dots, \beta_K)^T$.

$$\begin{aligned} \int f''(x)^2 dx &= \int \left(\sum_{i=1}^K \beta_i M_i''(x) \right) \left(\sum_{j=1}^K \beta_j M_j''(x) \right) dx = \sum_{i=1}^K \sum_{j=1}^K \beta_i \beta_j \int M_i''(x) M_j''(x) dx \\ &= [\beta_1 \dots \beta_K] \begin{bmatrix} \int M_1''(x) M_1''(x) dx & \dots & \int M_1''(x) M_K''(x) dx \\ \vdots & & \vdots \\ \int M_K''(x) M_1''(x) dx & \dots & \int M_K''(x) M_K''(x) dx \end{bmatrix} \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_K \end{bmatrix} = \beta^T \begin{bmatrix} \int M_1''(x) M_1''(x) dx & \dots & \int M_1''(x) M_K''(x) dx \\ \vdots & & \vdots \\ \int M_K''(x) M_1''(x) dx & \dots & \int M_K''(x) M_K''(x) dx \end{bmatrix} \beta \end{aligned}$$

+4 (2) [+4] Calculate $\int M_2''(t)^2 dt$,

where $M_2(t) = \frac{I(\xi_1 \leq t < \xi_2)}{2\Delta} \{7z_1(t)^3 - 18z_1(t)^2 + 12z_1(t)\} - \frac{I(\xi_2 \leq t < \xi_3)}{2\Delta} z_3(t)^3$ ($\frac{d}{dt} z_i(t) = \frac{1}{\Delta}$)

$\rightarrow M_2'(t) = \frac{I(\xi_1 \leq t < \xi_2)}{2\Delta^2} (21z_1(t)^2 - 36z_1(t) + 12) - 3 \frac{I(\xi_2 \leq t < \xi_3)}{2\Delta^2} z_3(t)^2$

$\rightarrow M_2''(t) = \frac{I(\xi_1 \leq t < \xi_2)}{2\Delta^3} (42z_1(t) - 36) - 3 \frac{I(\xi_2 \leq t < \xi_3)}{\Delta^3} z_3(t)$

$\Rightarrow \int M_2''(t)^2 dt = \int \frac{I(\xi_1 \leq t < \xi_2)}{4\Delta^6} (42z_1(t) - 36)^2 dt + 9 \int \frac{I(\xi_2 \leq t < \xi_3)}{\Delta^6} z_3^2(t) dt$
 $- 3 \int \frac{I(\xi_1 \leq t < \xi_2) I(\xi_2 \leq t < \xi_3)}{\Delta^6} z_3(t) (42z_1(t) - 36) dt$
 this term equal to zero $\because I(\xi_1 \leq t < \xi_2) I(\xi_2 \leq t < \xi_3) = 0 \forall t$

$= \frac{1}{4\Delta^6} \int_{\xi_1}^{\xi_2} (42z_1(t) - 36)^2 dt + \frac{9}{\Delta^6} \int_{\xi_2}^{\xi_3} z_3^2(t) dt$
 (Let $x = z_1(t) = \frac{t - \xi_1}{\Delta}$ and let $y = z_3(t) = \frac{t - \xi_3}{\Delta}$
 $\rightarrow dt = \Delta dx$ and $\rightarrow dt = \Delta dy$)

$= \frac{1}{4\Delta^6} \int_0^1 \Delta (42x - 36)^2 dx + \frac{9}{\Delta^6} \int_{-1}^0 \Delta y^2 dy$

$= \frac{1}{4\Delta^5} x \left. \frac{(42x - 36)^3}{3 \times 42} \right|_{x=0}^{x=1} + \frac{9}{\Delta^5} x \left. \frac{y^3}{3} \right|_{y=-1}^{y=0}$

$= \frac{1}{4\Delta^5} x \frac{6^3 + 36^3}{126} + \frac{9}{\Delta^5} x \frac{1}{3} = \frac{1}{4\Delta^5} \times 372 + \frac{3}{\Delta^5}$

$= \frac{93}{\Delta^5} + \frac{3}{\Delta^5} = \frac{96}{\Delta^5}$

+2 (3) [+2] Let Ω be a matrix with elements $\Omega_{jk} = \int M_j''(t)M_k''(t)dt$. Let $\{(Y_i, x_i), i=1, \dots, n\}$ be data.

Let $\hat{f}_\lambda(x) = \sum_{j=1}^K \hat{\beta}_{j\lambda} M_j(x)$ be a smoothing spline. Derive the formula of $\hat{\beta}_\lambda = (\hat{\beta}_{1\lambda}, \dots, \hat{\beta}_{K\lambda})^T$.

RSS with penalty of roughness = $RSS(\beta) = (y - X\beta)^T (y - X\beta) + \lambda \int f''(x)^2 dx$
 $\checkmark (y - X\beta)^T (y - X\beta) + \lambda \beta^T \Omega \beta$ where $y = \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}$, $\beta = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_K \end{bmatrix}$, $X = \begin{bmatrix} M_1(x_1) & \dots & M_K(x_1) \\ \vdots & & \vdots \\ M_1(x_n) & \dots & M_K(x_n) \end{bmatrix}$

$\Rightarrow \hat{\beta}_\lambda = \arg \min_{\beta} RSS(\beta) = \arg \min_{\beta} (y - X\beta)^T (y - X\beta) + \lambda \beta^T \Omega \beta$

That equivalent to $\frac{\partial}{\partial \beta} RSS(\beta) \Big|_{\hat{\beta}_\lambda} = -2X^T(y - X\hat{\beta}_\lambda) + 2\lambda \Omega \hat{\beta}_\lambda \stackrel{set}{=} 0$

$\Rightarrow (X^T X + \lambda \Omega) \hat{\beta}_\lambda = X^T y \Rightarrow \hat{\beta}_\lambda = (X^T X + \lambda \Omega)^{-1} X^T y$

+2 (4) [+2] Define $CV(\lambda)$.

$CV(\lambda) = \frac{1}{n} \sum_{i=1}^n \left(\frac{Y_i - \hat{f}_\lambda(x_i)}{1 - h_i} \right)^2$

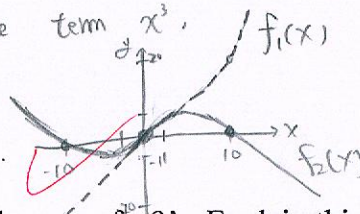
where h_i is the i -th diagonal element of $\checkmark X(X^T X + \lambda \Omega)^{-1} X^T$

+2 (5) [+2] The polynomial regression is sensitive to changes of β 's. Explain this by a figure.

Here compare these to function $\begin{cases} f_1(x) = x + \frac{1}{100} x^3 \\ f_2(x) = x - \frac{1}{100} x^3 \end{cases}$

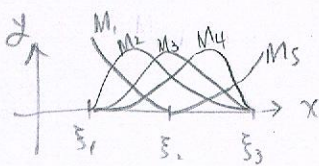
The coefficient $\frac{1}{100}$ can be treated as a small value, but it has large influence of the term x^3 .

X	-10	-1	0	1	10
f_1	-20	-1.01	0	1.01	20
f_2	0	-0.99	0	0.99	0



\rightarrow The performance of f_1 and f_2 are very different, this show that polynomial regression is sensitive to the changes of β 's.

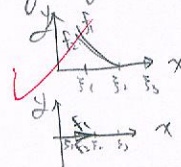
+2 (6) [+2] The cubic spline is not sensitive to changes of β 's. Explain this by a figure.



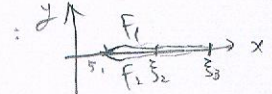
We see the figure marginally,

1. $\begin{cases} f_1(x) = 1.01 M_1(x) \\ f_2(x) = 0.99 M_1(x) \end{cases}$

2. $\begin{cases} f_1(x) = 0.01 M_1(x) \\ f_2(x) = -0.01 M_2(x) \end{cases}$



3. $\begin{cases} f_1(x) = 0.01 M_2(x) \\ f_2(x) = -0.01 M_2(x) \end{cases}$



etc.

we can see the small change of β_j 's does not have obviously influence on the different model marginally, and by the continuity of cubic spline, a very small difference of each β 's also no influence the result obviously (not sensitive to changes of β 's)